



## A new statistical approach for interpreting oceanic $f\text{CO}_2$ data



Hongjie Wang<sup>a</sup>, Xinping Hu<sup>a,\*</sup>, Blair Sterba-Boatwright<sup>b,\*</sup>

<sup>a</sup> Department of Physical and Environmental Sciences, Texas A&M University – Corpus Christi, Corpus Christi, TX 78412, United States

<sup>b</sup> Department of Mathematics & Statistics, Texas A&M University – Corpus Christi, Corpus Christi, TX 78412, United States

### ARTICLE INFO

#### Article history:

Received 4 November 2015

Received in revised form 18 May 2016

Accepted 23 May 2016

Available online 24 May 2016

#### Keywords:

Surface ocean  $f\text{CO}_2$

Ocean margin

Generalized Additive Mixed Modeling

Heteroscedasticity

### ABSTRACT

Surface ocean  $\text{CO}_2$  fugacity ( $f\text{CO}_2$ ) often exhibits large fluctuations and heterogeneity because of multiple controlling factors that pose a challenge for trend analysis, especially in the ocean margins. We propose a new statistical approach, Generalized Additive Mixed Modeling (GAMM), to interpret oceanic  $f\text{CO}_2$  data in two ocean margins (Japan and Europe) and two open ocean areas. The latter included areas near the Hawaii Ocean Time-series (HOT) and the Bermuda Atlantic Time-series (BATS). This method utilizes day of year, sea surface salinity (SSS), sea surface temperature (SST), and sampling date as predictors. Using this method, we were able to derive multidecadal  $f\text{CO}_2$  trends with both improved precision and greater robustness to data gaps compared to an existing deseasonalization method used in the open ocean. The  $f\text{CO}_2$  trend derived by our method for the Japanese margin (1992–2013), the European margin (1989–2014), and the open ocean near HOT (1983–2013) were  $2.1 \pm 0.6$ ,  $1.9 \pm 0.7$ , and  $2.0 \pm 0.5 \mu\text{atm year}^{-1}$  (mean  $\pm$  standard deviation of multiple  $1^\circ \times 1^\circ$  grids in margins and  $5^\circ \times 5^\circ$  grids in open ocean), respectively, and the  $f\text{CO}_2$  trends were all close to the atmosphere  $\text{CO}_2$  trend ( $1.7\text{--}1.9 \mu\text{atm year}^{-1}$ ). Our analysis produced generally smaller standard errors (paired  $t$ -test,  $p < 0.001$ ) than those obtained using the existing method based on the same dataset. In addition, our method was less sensitive to data gaps compared to this existing method. However, for regularly spaced  $f\text{CO}_2$  times-series data, for example, discrete bottle data collected in the BATS station in 1991–2011, this method was not advantageous over the existing method ( $1.9 \pm 0.2$  vs.  $2.0 \pm 0.2 \mu\text{atm year}^{-1}$ ). To test the broader applicability of this method, we compared  $f\text{CO}_2$  trends in the Southern Ocean derived using our method with those from a recently reported Markov Chain Monte Carlo method and found no significant difference between the two sets of values. Therefore, we recommend the application of our method in interpreting  $f\text{CO}_2$  data in different oceanic environments.

© 2016 Elsevier B.V. All rights reserved.

### 1. Introduction

The concentration of carbon dioxide ( $\text{CO}_2$ ) in the atmosphere has increased from  $\sim 280$  ppm in preindustrial era to  $\sim 399$  ppm (annual mean) in 2015 (<http://www.esrl.noaa.gov/gmd/ccgg/trends/>). If  $\text{CO}_2$  emission continues under the “business-as-usual” scenario, atmospheric  $\text{CO}_2$  level is expected to exceed 900 ppm by the end of this century (Collins et al., 2013).

The ocean serves as a natural sink ( $1.4\text{--}2.9 \text{ Pg C year}^{-1}$ ) for atmospheric  $\text{CO}_2$  (Khatiwala et al., 2009; Landschützer et al., 2014; Sabine et al., 2004; Takahashi et al., 2009). It takes up  $\text{CO}_2$  primarily through air-sea gas exchange, which is a function of both physical conditions (wind speed, salinity, and temperature) and the thermodynamic gradient between seawater and the atmosphere ( $\Delta f\text{CO}_2 = f\text{CO}_{2\_ocean} - f\text{CO}_{2\_air}$ ). Surface ocean  $\text{CO}_2$  fugacity ( $f\text{CO}_2$ ) is controlled by dissolved

inorganic carbon (DIC) concentration, total alkalinity (TA), sea surface salinity (SSS), and sea surface temperature (SST) (Takahashi et al., 1993). Both TA and DIC can also be affected by biogeochemical processes, including respiration, photosynthesis, and carbonate precipitation and dissolution (Zeebe and Wolf-Gladrow, 2001). Note in the literature both  $f\text{CO}_2$  and  $p\text{CO}_2$  (partial pressure of  $\text{CO}_2$ ) are often used interchangeably. Given that these two parameters are very close to each other ( $\Delta < 0.4\%$  (Zeebe and Wolf-Gladrow, 2001)) under normal seawater conditions, trend analysis based on either dataset would produce essentially the same results.

The ocean's uptake of  $\text{CO}_2$  has reduced atmospheric buildup of this greenhouse gas and damped associated climate changes. However,  $\text{CO}_2$  absorption has also reduced the saturation state of calcium carbonate, known as “ocean acidification” (Doney et al., 2009; Hoegh-Guldberg et al., 2007; Orr et al., 2005), which could hinder carbonate formation in marine organisms (e.g., corals, marine plankton, coralline algae, and shellfish) (Kleypas et al., 1999; Orr et al., 2005; Waldbusser et al., 2015).

The change rate of oceanic  $f\text{CO}_2$  relative to that in the atmosphere provides information on how the strength of the ocean as either a  $\text{CO}_2$

\* Corresponding authors.

E-mail addresses: [xinping.hu@tamucc.edu](mailto:xinping.hu@tamucc.edu) (X. Hu), [sterbaboatwright@tamucc.edu](mailto:sterbaboatwright@tamucc.edu) (B. Sterba-Boatwright).

sink or source evolves over time. For example, a region with an oceanic  $f\text{CO}_2$  increase rate higher than the atmospheric rate can be interpreted as a decreasing sink or an increasing source depending on the initial seawater  $f\text{CO}_2$  relative to atmospheric  $f\text{CO}_2$ . Conversely, an oceanic  $f\text{CO}_2$  increase rate less than the atmospheric rate can be interpreted as an increasing sink (or a decreasing source) (Landschützer et al., 2014; Landschützer et al., 2015; Lenton et al., 2012; Majkut et al., 2014). On the other hand, because of high primary productivity in the ocean margins (Liu et al., 2010), knowledge on the  $f\text{CO}_2$  changes there is essential for understanding its role in global carbon cycle.

Getting a more precise  $f\text{CO}_2$  trend is crucial to understanding the evolution of carbon sink/source in difference areas of the ocean. In the literature, however, there is no universal method in calculating the  $f\text{CO}_2$  trend in surface seawater. The simplest approach is a linear least square regression using observed data. For example, Dore et al. (2009) directly applied a linear regression using observed  $f\text{CO}_2$  data in waters at the Hawaii Ocean Time-series (HOT) and found that the trend in the surface ocean is  $1.9 \pm 0.2 \mu\text{atm year}^{-1}$ . To avoid biases caused by temporal weighting, some researchers chose the observed data in only selected months for their calculations. For example, Midorikawa et al. (2010) reported that the  $f\text{CO}_2$  trend in the western North Pacific (1983–2007) is  $1.58 \pm 0.12 \mu\text{atm year}^{-1}$  in winter and  $1.37 \pm 0.33 \mu\text{atm year}^{-1}$  in summer, and neither trend is significantly different from their respective air  $f\text{CO}_2$  trends ( $1.65 \pm 0.05 \mu\text{atm year}^{-1}$  in winter and  $1.54 \pm 0.08 \mu\text{atm year}^{-1}$  in summer). Similarly, to minimize the biological effects on data interpretation in the Iceland Sea, Olafsson et al. (2009) only selected the first 67 days (presumably with little biological production) of the sampled years (1985–2008) to calculate  $f\text{CO}_2$  trend, then with the aid of multivariate linear regression ( $y = a \times \text{time} + b \times \text{Temp} + c$ ), they found that the  $f\text{CO}_2$  trend is  $2.1 \pm 0.2 \mu\text{atm year}^{-1}$ .

The most commonly used approaches generally obtain linear trends after various deseasonalization techniques. For example, the  $f\text{CO}_2$  trend analysis from 1983 to 2010 in Bermuda Atlantic Time-series (BATS) was performed using both raw data and deseasonalized data (Bates et al., 2012). The rate is  $1.8 \pm 0.1 \mu\text{atm year}^{-1}$  using the deseasonalized data, but  $1.6 \pm 0.2 \mu\text{atm year}^{-1}$  using the raw data. This latter trend may be biased due to more frequent sampling in spring, thus Bates et al. (2012) recommended the deseasonalization method. To fit the seasonal cycle in the European Station for Time series in the ocean (ESTOC) near the Canary Islands, Santana-Casiano et al. (2007) used harmonic functions to decompose the time series into a trend, seasonal variations and errors, and serial correlation was modeled using a second-order autoregressive process. They found that the  $f\text{CO}_2$  trend from 1995 to 2004 is  $1.6 \pm 0.4 \mu\text{atm year}^{-1}$ . In another study, Schuster et al. (2009) fitted a harmonic function in the form of  $y = a + b \times t + c \times \cos(2\pi t + d)$  in the North Atlantic to calculate the  $f\text{CO}_2$  trend, where  $t$  is the year lapse since an arbitrarily defined reference year, and  $b$  is the  $f\text{CO}_2$  trend in  $\mu\text{atm year}^{-1}$ . Their results suggested that sea-surface  $f\text{CO}_2$  has closely followed atmospheric  $f\text{CO}_2$  in the subtropical regions. McKinley et al. (2011) and Fay and McKinley (2013) also adopted this harmonic function. They found that the  $f\text{CO}_2$  trend in the open ocean is sensitive to the chosen start and end years, resulting from climatic oscillations such as El Niño/Southern Oscillation, North Atlantic Oscillation. However, these oscillation signals would fade away as timescales increase (i.e., 25 years), and the  $f\text{CO}_2$  trend is parallel to the atmospheric trend (Fay and McKinley, 2013; McKinley et al., 2011).

Takahashi et al. (2009) proposed a simple yet effective deseasonalization method and used it to obtain global surface  $p\text{CO}_2$  climatology. In their study, all historical data from regularly spaced “grids” (latitude  $\times$  longitude of either  $4^\circ \times 5^\circ$  or  $5^\circ \times 10^\circ$ ) were used to obtain the rate in each grid. First, seasonal changes were calculated on the basis of the monthly mean values computed from a 4-year subsample of the entire time series. Then  $p\text{CO}_2$  values for months with no measurements were estimated by a linear interpolation using two adjacent

monthly means. The difference between a monthly mean and the annual mean represents the correction to be applied to deseasonalize the monthly mean. Finally, the deseasonalized monthly mean values are regressed against time (year) using least square method to obtain the mean rate of change (Lenton et al., 2012; Takahashi et al., 2009). Using their method, surface water  $p\text{CO}_2$  in the North Atlantic, North and South Pacific and Southern Oceans increases at a mean rate of  $1.5 \mu\text{atm year}^{-1}$  from 1970 to 2007 (Takahashi et al., 2009). Overall, different independent studies suggested that surface  $f\text{CO}_2$  trend in the open ocean has increased more or less the same as the atmospheric  $f\text{CO}_2$  has. These trends indicate that the driving force for air-sea  $\text{CO}_2$  gas exchange has not changed significantly over the last three decades (Bates et al., 2012).

Recently, Majkut et al. (2014) developed a Markov Chain Monte Carlo (MCMC) method to calculate long-time  $p\text{CO}_2$  trends in the open ocean. They found that the  $p\text{CO}_2$  trend in the Southern Ocean is  $1.4 \pm 0.5 \mu\text{atm year}^{-1}$  (based on the Lamont-Doherty Earth Observatory LDEO V2010 database) in the 1995–2008 period, and they suggested a global increase in the  $\text{CO}_2$  uptake of  $0.4 \pm 0.1 \text{Pg C year}^{-1} \text{decade}^{-1}$ , because surface  $p\text{CO}_2$  is increasing more slowly than the atmospheric value (but not significantly different from the latter). This method resulted a smaller value than linear regression method ( $2.2 \pm 0.2 \mu\text{atm year}^{-1}$ ) reported in Lenton et al. (2012) based on the same dataset. Using a neural network approach, Landschützer et al. (2015) reported the weakening carbon sink trend in the Southern ocean (south of  $35^\circ\text{S}$ ) stopped around 2002 and the carbon sink there increasing from  $\sim 0.6 \text{Pg C year}^{-1}$  in 2002 to  $\sim 1.2 \text{Pg C year}^{-1}$  in 2011.

One of the assumptions required for inference with regression models is that residuals have constant variance throughout the range of the predictors and the fitted values (“homoscedasticity”). Unfortunately, in most cases this statistical test was not performed or explicitly demonstrated, even in the sophisticated models published recently, for example, the MCMC approach (Majkut et al., 2014) or the neural network approach (Landschützer et al., 2013; Landschützer et al., 2015). Given the spatial heterogeneity of biogeochemical reactions in marginal areas, which are heavily modulated by terrestrial influences, biological activities, and physical processes (such as upwelling) (Liu et al., 2010), it is unknown whether the above deseasonalization method can be used in these areas, where the cyclic (or seasonal) behavior may not be as stable as that in the open sea (Cai, 2011; Liu et al., 2010). The assumption of homoscedasticity is not necessarily met because of heterogeneity of the marginal ocean and large  $f\text{CO}_2$  fluctuations. Violation of this assumption suggests that linear least squares may not necessarily result in the best unbiased estimates of model parameters. Therefore, the regression analysis obtained from heteroscedastic variables could produce inaccurate standard errors, too large for some values and too small in others, and potentially biased regression coefficients.

In this work, we used a statistical approach that can fit the seasonal cycle more precisely and reduce the impact of heteroscedasticity on the calculated trend, and we compared results from our method with those from the Takahashi method (hereafter T0) that effectively generated open ocean  $p\text{CO}_2$  climatology (Takahashi et al., 2009). We also compared  $f\text{CO}_2$  trend derived using our method with recently reported results in the Southern Ocean using the MCMC method to test a broader applicability of our method.

## 2. Method

### 2.1. Data

We chose two marginal areas and two open ocean areas in this study. The two marginal areas include the northwestern North Pacific margin (east of Japan,  $28\text{--}45^\circ\text{N}$ ,  $130\text{--}150^\circ\text{E}$ ) and northeastern North Atlantic offshore Europe ( $44\text{--}50^\circ\text{N}$ ,  $3\text{--}14^\circ\text{W}$ ). The two open ocean areas are near the two time series stations, HOT ( $10\text{--}30^\circ\text{N}$ ,  $140\text{--}165^\circ\text{W}$ ) and BATS ( $31^\circ 40'\text{N}$ ,  $64^\circ 10'\text{W}$ ), respectively. The grid size for our model

Download English Version:

<https://daneshyari.com/en/article/1260904>

Download Persian Version:

<https://daneshyari.com/article/1260904>

[Daneshyari.com](https://daneshyari.com)