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## Short communication: Multi-trait estimation of genetic parameters for milk protein composition in the Danish Holstein

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### ABSTRACT

Genetic parameters were estimated for the major milk proteins using bivariate and multi-trait models based on genomic relationships between animals. The analyses included, apart from total protein percentage,  $\alpha_{S1}$ -casein (CN),  $\alpha_{S2}$ -CN,  $\beta$ -CN,  $\kappa$ -CN,  $\alpha$ -lactalbumin, and  $\beta$ -lactoglobulin, as well as the posttranslational sub-forms of glycosylated  $\kappa$ -CN and  $\alpha_{s1}$ -CN-8P (phosphorylated). Standard errors of the estimates were used to compare the models. In total, 650 Danish Holstein cows across 4 parities and days in milk ranging from 9 to 481 d were selected from 21 herds. The multi-trait model generally resulted in lower standard errors of heritability estimates, suggesting that genetic parameters can be estimated with high accuracy using multitrait analyses with genomic relationships for scarcely recorded traits. The heritability estimates from the multi-trait model ranged from low (0.05 for  $\beta$ -CN) to high (0.78 for  $\kappa$ -CN). Genetic correlations between the milk proteins and the total milk protein percentage were generally low, suggesting the possibility to alter protein composition through selective breeding with little effect on total milk protein percentage.

Key words: genetic parameter, milk protein, multitrait model, genomic relationship

#### **Short Communication**

Milk protein composition plays an important role in the technological properties of milk (Ikonen et al., 1999; Bittante et al., 2012). Changes in relative concentrations of individual milk proteins have a major effect on milk coagulation properties (Bonfatti et al., 2011) and coagulation ability of milk is essential in cheese making (Cassandro et al., 2008). The major milk proteins include  $\alpha_{s_1}$ -CN,  $\alpha_{s_2}$ -CN,  $\beta$ -CN,  $\kappa$ -CN,  $\alpha$ -LA, and  $\beta$ -LG. In addition, several posttranslational modifications of these proteins exist in milk.

Previous studies have shown that considerable genetic variation exists in the composition of milk protein (Bobe et al., 1999; Schopen et al., 2009), presenting the opportunity to alter milk protein composition through selective breeding. Reliable estimates of genetic parameters, including heritability and genetic covariances, are crucial to evaluate the potential for breeding. Quantifying specific milk proteins requires specialized and costly equipment, making it difficult and expensive to measure the traits. As a result, sufficient phenotypic data are not available for reliable estimation of genetic parameters. One effective strategy to deal with such scarcely recorded traits could be implementation of multi-trait models that take advantage of information from correlated traits (Calus and Veerkamp, 2011).

Generally, only a few studies have previously estimated genetic parameters for specific milk proteins (Schopen et al., 2009; Bonfatti et al., 2011) and their posttranslational sub-forms (Bijl et al., 2014). More importantly, none of the previous studies have estimated genetic parameters for milk protein profile using multitrait analyses.

In this study, we estimated genetic parameters for the major milk proteins ( $\alpha_{S1}$ -CN,  $\alpha_{S2}$ -CN,  $\beta$ -CN,  $\kappa$ -CN,  $\alpha$ -LA and  $\beta$ -LG), the posttranslational sub-forms (glycosylated  $\kappa$ -CN and  $\alpha_{S1}$ -CN-8P, where P = phosphorylated serine), as well as protein percentage using bivariate and multi-trait models with genomic relationships between animals and compared standard errors of the estimated genetic parameters.

Morning milk samples were obtained from 650 cows from 21 herds in Denmark. The cows were in different stages of lactation (d 9 to 481 in milk) and parity 1 to 4. The liquid chromatography/electrospray ionizationmass spectrometry (LC/ESI-MS) methods were used to profile the milk proteins. Details on screening of samples and quantification of milk proteins were previously described by Jensen et al. (2012).

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Of the total cows, 372 were genotyped using the BovineHD Illumina BeadChip. The remaining 278 cows were genotyped with the BovineSNP50 beadchip. Genomic DNA was extracted from ear tissue. Genotypes were subsequently imputed to full sequence in a 2-step procedure. The 278 cows genotyped with the BovineSNP50 chip were first imputed to the BovineHD (777k) level using a multi-breed reference of 3,383 animals including the 372 HD genotyped cows used in this study. The true and imputed HD data for the 2 cow groups were then merged and imputed to the wholegenome sequence level using a multi-breed reference of 1,228 animals from the "1000 bull genomes" project (http://www.1000bullgenomes.com/) and data from Aarhus University using IMPUTE2 v2.3.1 (Howie et al., 2011).

The genomic relationship matrix was calculated as described by the first method presented in VanRaden (2008). In total, 3.7 million SNP markers spread over BTA1 to BTA29 were included to calculate the **G** matrix.

The REML approach in DMU was used to estimate genetic parameters and variance components (Madsen and Jensen, 2010). Bivariate and multi-trait analyses were performed and compared using standard errors for the estimated heritability.

The general model used was

$$y_{ijkl} = \mu + parity_i + herd_j + b_1 \times DIM_k + b_2$$
$$\times exp^{-0.05 \times DIM_k} + animal_l + e_{iikl}, \qquad [1]$$

where  $y_{ijkl}$  was the observation of animal l, in parity i and herd j;  $\mu$  was the fixed mean effect;  $b_1$  was the regression coefficient for  $\text{DIM}_k$ ; and  $\text{DIM}_k$  was a covariate describing the effect of days k in milk. Wilmink adjustment (exp<sup>-0.05×DIM</sup>) was used for DIM,  $b_2$  was the re-

gression coefficient for the Wilmink adjustment; animal<sub>l</sub> was the random additive genetic effect based on **G** of animal *l* with distribution  $N(0, \mathbf{G}\sigma_a^2)$ , and  $\mathbf{e}_{ijkl}$  was the random residual effect, which was assumed to be normally distributed with  $e \sim N(0, \mathbf{I}\sigma_e^2)$ , where **G** is the genomic relationship matrix, **I** was the identity matrix,  $\sigma_a^2$  was the genetic variation, and  $\sigma_e^2$  was the residual variation.

The bivariate analyses were run for each milk protein analyzed in combination with protein percentage. For the multi-trait analysis, all 9 traits were fitted simultaneously. Correlations between traits were based on the multi-trait analyses.

Table 1 summarizes the descriptive statistics for the milk protein profile and the total milk protein percentage. Mean protein content in the sampled milk was 3.38%. The major proteins ( $\alpha_{S1}$ -CN,  $\alpha_{S2}$ -CN,  $\beta$ -CN,  $\kappa$ -CN,  $\alpha$ -LA, and  $\beta$ -LG) made up 83% of the total milk protein fraction. The caseins constituted 72.3% of the total protein, with  $\beta$ -CN and  $\alpha_{S1}$ -CN alone contributing to 34.1 and 26.8% of the total milk protein, respectively. The whey proteins constituted 10.8% of the total protein.

The  $\alpha_{S1}$ -CN-8P accounted for 19.2% of the total milk protein and 71.6% of the  $\alpha_{S1}$ -CN fraction of the total protein percentage. This was comparable to previous findings, in which  $\alpha_{S1}$ -CN-8P accounted for 21.3% of the total protein (Bijl et al., 2014) and 74% of the  $\alpha_{S1}$ -CN (Heck et al., 2008) in the Dutch Holstein population.

Heritability values and standard errors of estimation from the bivariate and multi-trait models are given in Table 2. Generally, the heritability estimates for the milk proteins were moderate to high except for  $\beta$ -CN, which had the lowest estimates (0.01–0.05). Glycosylated  $\kappa$ -CN (0.44) and  $\alpha_{s2}$ -CN (0.36) had moderate

| Protein or fraction <sup>2</sup> | Mean $(\%)$ | CV (%) | 5% quantile | 95% quantile |
|----------------------------------|-------------|--------|-------------|--------------|
| α <sub>S1</sub> -CN              | 26.8        | 9      | 25.5        | 28.1         |
| $\alpha_{S1}$ -CN-8P             | 19.2        | 11     | 17.7        | 20.7         |
| as2-CN                           | 5.3         | 20     | 4.5         | 5.9          |
| β-CN                             | 34.1        | 10     | 31.5        | 36.7         |
| κ-CN                             | 6.1         | 18     | 5.3         | 6.9          |
| Glycosylated $\kappa$ -CN        | 1.7         | 47     | 1.2         | 2.0          |
| α-LA                             | 3.3         | 19     | 2.9         | 3.6          |
| β-LG                             | 7.5         | 21     | 6.5         | 8.4          |
| Total protein (%)                | 3.38        | 9      | 3.20        | 3.55         |

Table 1. Descriptive statistics<sup>1</sup> of milk protein profile and the total milk protein percentage

 $^{1}Mean = phenotypic mean of the trait.$ 

<sup>2</sup>Protein composition was expressed as percentage fractions of the total milk protein percentage (wt/wt); total protein was expressed as percentage (%) of the total milk yield; individual proteins comprise only the peaks identified as intact proteins and isoforms; that is,  $\alpha_{SI}$ -CN (comprises  $\alpha_{SI}$ -CN 8P + 9P),  $\alpha_{S2}$ -CN (comprises  $\alpha_{S2}$ -CN 11P + 12P),  $\beta$ -CN (comprises  $\beta$ -CN 4P + 5P), and  $\kappa$ -CN (comprises  $\kappa$ -CN G + 1P), where P = phosphorylated group.

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