

# Tests for assessment of agreement using probability criteria

Pankaj K. Choudhary<sup>a,\*</sup>, H.N. Nagaraja<sup>b</sup>

<sup>a</sup>Department of Mathematical Sciences, University of Texas at Dallas, Richardson, TX 75083-0688, USA

<sup>b</sup>Department of Statistics, Ohio State University, Columbus, OH 43210-1247, USA

Received 18 May 2004; accepted 11 November 2005

Available online 19 January 2006

---

## Abstract

For the assessment of agreement using probability criteria, we obtain an exact test, and for sample sizes exceeding 30, we give a bootstrap- $t$  test that is remarkably accurate. We show that for assessing agreement, the total deviation index approach of Lin [2000. Total deviation index for measuring individual agreement with applications in laboratory performance and bioequivalence. *Statist. Med.* 19, 255–270] is not consistent and may not preserve its asymptotic nominal level, and that the coverage probability approach of Lin et al. [2002. Statistical methods in assessing agreement: models, issues and tools. *J. Amer. Statist. Assoc.* 97, 257–270] is overly conservative for moderate sample sizes. We also show that the nearly unbiased test of Wang and Hwang [2001. A nearly unbiased test for individual bioequivalence problems using probability criteria. *J. Statist. Plann. Inference* 99, 41–58] may be liberal for large sample sizes, and suggest a minor modification that gives numerically equivalent approximation to the exact test for sample sizes 30 or less. We present a simple and accurate sample size formula for planning studies on assessing agreement, and illustrate our methodology with a real data set from the literature.

© 2006 Elsevier B.V. All rights reserved.

*Keywords:* Bootstrap; Concordance correlation; Coverage probability; Limits of agreement; Total deviation index; Tolerance interval

---

## 1. Introduction

Suppose a paired sample of reference and test measurements,  $(X, Y)$ , are available on  $n$  randomly chosen subjects from a population of interest. Generally the  $Y$ 's are cheaper, faster, easier or less invasive to obtain than the  $X$ 's. The question of our concern is: "Are  $X$  and  $Y$  close enough so that they can be used interchangeably?" This comparison is the goal of method comparison studies. We focus on a test of hypotheses approach for this problem that infers *satisfactory agreement* between  $Y$  and  $X$  when the difference  $D = Y - X$  lies within an acceptable margin with a sufficiently high probability. We will assume that  $D$  follows a  $N(\mu, \sigma^2)$  distribution. Let  $F(\cdot)$  and  $\Phi(\cdot)$  denote the c.d.f.'s of  $|D|$  and a  $N(0, 1)$  distribution, respectively.

There are two measures of agreement based on the probability criteria. The first is the  $p_0$ th percentile of  $|D|$ , say  $Q(p_0)$ , where  $p_0 (> 0.5)$  is a specified large probability (usually  $\geq 0.80$ ). It was introduced by Lin (2000) who called it the *total deviation index* (TDI). Its small value indicates a good agreement between  $(X, Y)$ . The TDI can be

---

\* Corresponding author. Tel.: +1 972 883 4436; fax: +1 972 883 6622.

E-mail addresses: [pankaj@utdallas.edu](mailto:pankaj@utdallas.edu) (P.K. Choudhary), [hnn@stat.ohio-state.edu](mailto:hnn@stat.ohio-state.edu) (H.N. Nagaraja).

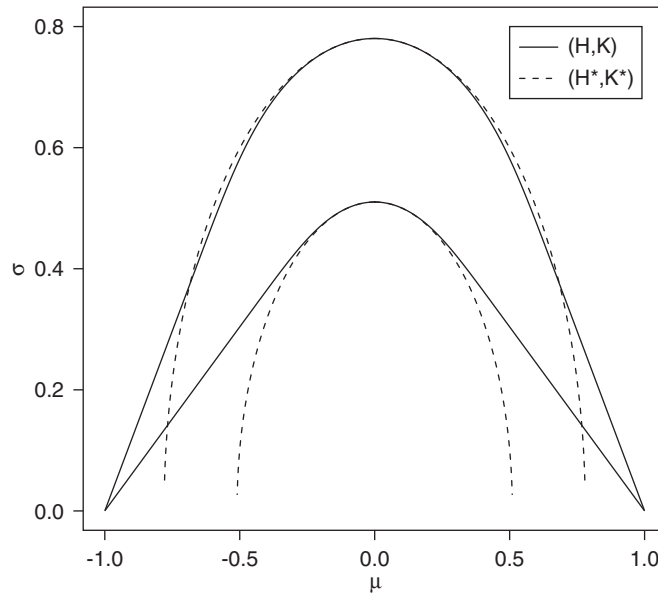


Fig. 1. The regions under the hypotheses (H, K) of our interest, given by (4) or (5), and those under (H\*, K\*) of Lin (2000), given by (7). The top two curves represent the boundaries of the two null regions for  $p_0 = 0.80$ , and the bottom two for  $p_0 = 0.95$ . The alternative regions lie under the respective curves. Here we have taken  $\delta_0 = 1.0$ , so that the x- and the y-axes actually represent  $\mu/\delta_0$  and  $\sigma/\delta_0$ , respectively.

expressed as,

$$Q(p_0) = F^{-1}(p_0) = \sigma\{\chi_1^2(p_0, \mu^2/\sigma^2)\}^{1/2}, \tag{1}$$

where  $\chi_1^2(p_0, \Delta)$  is the  $p_0$ th percentile of a  $\chi^2$ -distribution with single degree of freedom and non-centrality parameter  $\Delta$ .

The second measure, introduced by Lin et al. (2002), is the coverage probability (CP) of the interval  $[-\delta_0, \delta_0]$ , where a difference under  $\pm\delta_0$  is considered practically equivalent to zero. There is no loss of generality in taking this interval to be symmetric around zero as it can be achieved by a location shift. Letting,

$$d_l = (-\delta_0 - \mu)/\sigma, \quad d_u = (\delta_0 - \mu)/\sigma, \tag{2}$$

the CP can be expressed as

$$F(\delta_0) = \Phi(d_u) - \Phi(d_l). \tag{3}$$

A high value of  $F(\delta_0)$  implies a good agreement between the methods.

For specified  $(p_0, \delta_0)$ ,  $F(\delta_0) \leq p_0 \iff Q(p_0) \geq \delta_0$ . Consequently, for assessing agreement one can test either the hypotheses

$$H: Q(p_0) \geq \delta_0 \quad \text{vs.} \quad K: Q(p_0) < \delta_0, \tag{4}$$

or

$$H: F(\delta_0) \leq p_0 \quad \text{vs.} \quad K: F(\delta_0) > p_0, \tag{5}$$

and infer satisfactory agreement if H is rejected.

Let  $\Theta = \{(\mu, \sigma) : |\mu| < \infty, 0 < \sigma < \infty\}$  be the parameter space. Also let  $\Theta_H$  and  $\Theta_K$  be the subsets of  $\Theta$  representing the regions under H and K, respectively, and  $\Theta_B$  be the boundary of H. These regions can be visualized through the solid curves in Fig. 1 for  $p_0 = 0.80, 0.95$ . Note that they are symmetric in  $\mu$  about zero.

Lin (2000) suggests a large sample test for the hypotheses (4) and we refer to it as the TDI test. Lin et al. (2002) suggest a large sample test for (5), and we call it the CP test. They also conclude that these tests are more powerful for inferring

Download English Version:

<https://daneshyari.com/en/article/1148727>

Download Persian Version:

<https://daneshyari.com/article/1148727>

[Daneshyari.com](https://daneshyari.com)