



# Distribution of order statistics from selected subsets of concomitants

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## ABSTRACT

For a random sample of size  $n$  from an absolutely continuous bivariate population  $(X, Y)$ , let  $X_{i:n}$  be the  $i$ th  $X$ -order statistic and  $Y_{[i:n]}$  be its concomitant. We study the joint distribution of  $(V_{s:m}, W_{t:n-m})$ , where  $V_{s:m}$  is the  $s$ th order statistic of the upper subset  $\{Y_{[i:n]}, i=n-m+1, \dots, n\}$ , and  $W_{t:n-m}$  is the  $t$ th order statistic of the lower subset  $\{Y_{[j:n]}, j=1, \dots, n-m\}$  of concomitants. When  $m=\lceil np_0 \rceil$ ,  $s=\lceil mp_1 \rceil$ , and  $t=\lceil (n-m)p_2 \rceil$ ,  $0 < p_i < 1, i=0,1,2$ , and  $n \rightarrow \infty$ , we show that the joint distribution is asymptotically bivariate normal and establish the rate of convergence. We propose second order approximations to the joint and marginal distributions with significantly better performance for the bivariate normal and Farlie–Gumbel bivariate exponential parents, even for moderate sample sizes. We discuss implications of our findings to data-snooping and selection problems.

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## 1. Introduction

Suppose  $(X_i, Y_i), i=1, \dots, n$ , is a random sample from an absolutely continuous bivariate population  $(X, Y)$ . If we order the sample by the  $X$ -variate, and obtain the order statistics,  $X_{1:n} < X_{2:n} < \dots < X_{n:n}$ , for the  $X$  sample, then the  $Y$ -variate associated with the  $r$ th order statistic  $X_{r:n}$  is called the *concomitant of the  $r$ th order statistic*, and is denoted by  $Y_{[r:n]}$ . The term *concomitant of order statistic* was introduced by David (1973) and independently Bhattacharya (1974) used the term *induced order statistic*. Concomitants of order statistics arise on several occasions. For a comprehensive review of the area see David and Nagaraja (1998); Sections 9.8 and 11.7 of David and Nagaraja (2003) contain a brief review. Recently He and Nagaraja (2009a) have proposed an efficient estimator of the correlation coefficient of a bivariate normal (BVN) distribution based on linear functions of the complete set of  $n$  concomitants and their order statistics. He and Nagaraja (2009b) have obtained the joint distribution of a single pair  $(Y_{i:n}, Y_{[j:n]})$ ,  $i, j=1, \dots, n$ , extracted from all of the  $n$  concomitants and used it to determine the expected cost of mismatch in broken bivariate samples.

In this paper we study the joint distribution of  $(V_{s:m}, W_{t:n-m})$ , where  $V_{s:m}$  is the  $s$ th order statistic of the upper subset  $\{Y_{[i:n]}, i=n-m+1, \dots, n\}$ , and  $W_{t:n-m}$  is the  $t$ th order statistic of the lower subset  $\{Y_{[j:n]}, j=1, \dots, n-m\}$  of concomitants. Nagaraja and David (1994) discussed the finite-sample and asymptotic distributions of the upper extreme order statistic  $V_{m:m}$  and Joshi and Nagaraja (1995) studied the joint distribution of extremes,  $(V_{m:m}, W_{n-m:n-m})$ . Chu et al. (1999) investigated the asymptotic distributions of extremes from central subsets of concomitants. Here we obtain results for central order statistics from upper and lower subsets, and discuss some interesting findings.

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We obtain in Section 2 the exact joint distribution of  $(V_{s:m}, W_{t:n-m})$  for the finite sample case using a conditioning argument. In Section 3 the asymptotic distribution of  $(V_{s:m}, W_{t:n-m})$  is obtained under appropriate regularity conditions assuming  $m = \lceil np_0 \rceil$ ,  $s = \lceil mp_1 \rceil$ , and  $t = \lceil (n-m)p_2 \rceil$ ,  $0 < p_i < 1, i = 0, 1, 2$ , as  $n \rightarrow \infty$ . A result on the rate of convergence is also provided. In Section 4 we propose a higher order approximation to these distributions. In Section 5, we illustrate our results with BVN and Farlie–Gumbel bivariate exponential (FGBVE) examples. For both distributions, we show that the second order approximation given in Section 4 performs markedly superior to the normal approximation even for moderate sample sizes. Section 6 discusses some applications of our results to data snooping and selection problems, and provides some generalizations of our results.

Let  $F(x, y)$  be the joint cdf of  $(X, Y)$ , and  $f(x, y)$  be its joint pdf;  $f_X(x)$  be the marginal pdf of  $X$ , and  $f_{Y|X}(y|x)$  be the conditional pdf of  $Y$  given  $X$ . Let  $F_1(y|x)(f_1(y|x))$  and  $F_2(y|x)(f_2(y|x))$  be the conditional cdf (pdf) of  $Y$  given  $X > x$  and  $X < x$ , respectively. For  $0 < p < 1$ , and a univariate cdf  $G$ , let  $G^{-1}(p)$  be the  $p$ th quantile. A bivariate normal distribution with mean  $\mu$  and covariance matrix  $\Sigma$  is denoted by  $\text{BVN}(\mu, \Sigma)$ . A standard normal pdf is denoted by  $\phi$ , cdf by  $\Phi$ , and a normal distribution with mean  $\mu$  and variance  $\sigma^2$  is denoted by  $N(\mu, \sigma^2)$ .

## 2. Finite-sample distribution of $(V_{s:m}, W_{t:n-m})$

We can derive the joint cdf of  $(V_{s:m}, W_{t:n-m})$  by conditioning on the value of  $X_{n-m:n}$  and using the following result established by Kaufmann and Reiss (1992) (Theorem 2):

**Lemma 1.** Given  $X_{n-m:n} = x$ ,  $V_{s:m}$  has the same distribution as the  $s$ th order statistic of a random sample of size  $m$  from the cdf  $F_1(\cdot|x)$ , and  $W_{t:n-m}$  has the same distribution as the  $t$ th order statistic of the sample consisting of  $n-m$  independent observations, of which  $n-m-1$  are from the cdf  $F_2(\cdot|x)$  and the remaining one is from the cdf  $F_{Y|X}(\cdot|x)$ . Moreover,  $V_{s:m}$  and  $W_{t:n-m}$  are conditionally independent given  $X_{n-m:n} = x$ .

From Lemma 1, we have

$$\begin{aligned} F_{V_{s:m}, W_{t:n-m}}(v, w) &= \mathbf{P}(V_{s:m} \leq v, W_{t:n-m} \leq w) \\ &= \int \mathbf{P}(V_{s:m} \leq v | X_{n-m:n} = x) \mathbf{P}(W_{t:n-m} \leq w | X_{n-m:n} = x) dF_{X_{n-m:n}}(x), \end{aligned}$$

where

$$\mathbf{P}(V_{s:m} \leq v | X_{n-m:n} = x) = \sum_{i=s}^m \binom{m}{i} [F_1(v|x)]^i [1 - F_1(v|x)]^{m-i}$$

and

$$\begin{aligned} \mathbf{P}(W_{t:n-m} \leq w | X_{n-m:n} = x) &= \sum_{j=t}^{n-m-1} \binom{n-m-1}{j} [F_2(w|x)]^j [1 - F_2(w|x)]^{n-m-1-j} + \binom{n-m-1}{t-1} [F_2(w|x)]^{t-1} [1 - F_2(w|x)]^{n-m-t} F_{Y|X}(w|x). \end{aligned} \quad (1)$$

**Remarks.** 1. The conditional pdfs  $f_1$  and  $f_2$  are given by

$$f_1(y|x) = \frac{\int_x^\infty f(u, y) du}{1 - F_X(x)}; \quad f_2(y|x) = \frac{\int_{-\infty}^x f(u, y) du}{F_X(x)}. \quad (2)$$

2. Note that the conditional cdf in (1) is just the cdf of an order statistic when there is a single outlier (Arnold and Balakrishnan, 1989, p. 109).

## 3. Asymptotic joint normality

### 3.1. Limit distribution of $(V_{s:m}, W_{t:n-m})$

**Theorem 1.** Suppose  $m = \lceil np_0 \rceil$ ,  $s = \lceil mp_1 \rceil$ , and  $t = \lceil (n-m)p_2 \rceil$ ,  $0 < p_i < 1, i = 0, 1, 2$ , as  $n \rightarrow \infty$ . Let  $x_0 = F_X^{-1}(q_0)$  with  $q_0 = 1 - p_0$ , and  $h_i(x) = F_i^{-1}(p_i|x)$ ,  $i = 1, 2$ , denote the  $p_i$ th quantile of the conditional cdf  $F_i(\cdot|x)$  when viewed as a function of  $x$ . We assume that  $h_i(x)$  has first order derivative at  $x_0$  and that  $f(x, y) > 0$  in a neighborhood of  $(x_0, h_i(x_0))$ . Define  $\sigma_0 = \sqrt{p_0 q_0} / f_X(x_0)$ , and let  $a_i = h_i(x_0)$ ,  $c_i = \sigma_0 h'_i(x_0)$ ,  $i = 1, 2$ , and

$$b_1 = \frac{\sqrt{p_1 q_1}}{\sqrt{p_0}} f_1(a_1|x_0), \quad b_2 = \frac{\sqrt{p_2 q_2}}{\sqrt{q_0}} f_2(a_2|x_0).$$

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