Contents lists available at ScienceDirect

Statistics and Probability Letters

journal homepage: www.elsevier.com/locate/stapro

Competitive estimation of the extreme value index*

M. Ivette Gomes^{a,b,*}, Lígia Henriques-Rodrigues^{c,b}

^a Universidade de Lisboa, Faculdade de Ciências (DEIO), Portugal

^b CEAUL, Portugal

^c Universidade de São Paulo, IME, Brazil

ARTICLE INFO

Article history: Received 4 May 2016 Accepted 14 May 2016 Available online 25 May 2016

Keywords: Heavy tails PORT methodology Statistical extreme value theory

ABSTRACT

The mean-of-order-p (MO_p) extreme value index (EVI) estimators are based on Hölder's mean of an adequate set of statistics, and generalize the classical Hill EVI-estimators, associated with p = 0. Such a class of estimators, dependent on the tuning parameter $p \in \mathbb{R}$, has revealed to be highly flexible, but it is not invariant for changes in location. To make the MO_p location-invariant, it is thus sensible to use the *peaks over a random threshold* (PORT) methodology, based upon the excesses over an adequate ascending order statistic. In this article, apart from an asymptotic comparison at optimal levels of the optimal MO_p class and some competitive EVI-estimators, like a Pareto probability weighted moments EVI-estimator, a few details on PORT classes of EVI-estimators are provided, enhancing their high efficiency both asymptotically and for finite samples.

© 2016 Elsevier B.V. All rights reserved.

1. Introduction

Given a sample of *n* random variables (RVs), $\underline{X}_n := (X_1, \ldots, X_n)$, either independent, identically distributed or possible weakly dependent and stationary from a *cumulative distribution function* (CDF) *F*, let us denote by $X_{1:n} \le \cdots \le X_{n:n}$ the associated ascending order statistics. Let us further assume that there exist sequences of real constants, $\{a_n > 0\}$ and $\{b_n \in \mathbb{R}\}$ such that the maximum, linearly normalized, i.e. $(X_{n:n} - b_n) / a_n$, converges in distribution to a non-degenerate RV. Then, the limiting CDF is compulsory of the type of an *extreme value* (EV) CDF, dependent upon a generalized shape parameter $\xi \in \mathbb{R}$, and with the functional form

$$EV_{\xi}(x) = \begin{cases} \exp(-(1+\xi x)^{-1/\xi}), \ 1+\xi x > 0 & \text{if } \xi \neq 0, \\ \exp(-\exp(-x)), \ x \in \mathbb{R} & \text{if } \xi = 0. \end{cases}$$
(1.1)

It is then said that *F* belongs to the max-domain of attraction of EV_{ξ} , in (1.1), and the notation $F \in \mathcal{D}_{\mathcal{M}}(EV_{\xi})$ is commonly used in the field of *extreme value theory* (EVT). The parameter ξ , the so-called *extreme value index* (EVI), is the primary parameter of extreme events. The EVI measures the heaviness of the *right tail-function* (RTF), $\overline{F}(x) := 1 - F(x)$, and the heavier the right-tail the larger the EVI is. For Paretian-type RTFs ($\xi > 0$), i.e. the often called heavy RTFs, the most popular EVI-estimator is the Hill (H) EVI-estimator (Hill, 1975), the average of the *k* log-excesses,

 $V_{ik} := \ln X_{n-i+1:n} - \ln X_{n-k:n}, \quad 1 \le i \le k < n,$

http://dx.doi.org/10.1016/j.spl.2016.05.012 0167-7152/© 2016 Elsevier B.V. All rights reserved.







[🌣] Research partially supported by FCT—Fundação para a Ciência e a Tecnologia, project UID/MAT/00006/2013 (CEA/UL).

^{*} Correspondence to: Faculdade de Ciências de Lisboa, Bloco C6, Piso 4, 1749-016 Lisboa, Portugal. E-mail address: ivette.gomes@fc.ul.pt (M.I. Gomes).

being thus the logarithm of the geometric mean (or mean-of-order-0) of

$$U_{ik} := X_{n-i+1:n} / X_{n-k:n}, \quad 1 \le i \le k < n,$$
(1.2)

i.e.

$$\hat{\xi}_{k}^{\mathsf{H}} = \hat{\xi}_{k}^{\mathsf{H}}(\underline{\mathbf{X}}_{n}) \coloneqq \frac{1}{k} \sum_{i=1}^{k} V_{ik} = \ln \left(\prod_{i=1}^{k} \frac{X_{n-i+1:n}}{X_{n-k:n}} \right)^{1/k}.$$
(1.3)

Brilhante et al. (2013) considered as basic statistics the mean-of-order-p (MO_p) of U_{ik} , in (1.2), for $p \ge 0$. More generally, Gomes and Caeiro (2014) considered the same statistics for $p \in \mathbb{R}$, i.e.

$$M_{p}(k) = \begin{cases} \left(\frac{1}{k}\sum_{i=1}^{k}U_{ik}^{p}\right)^{1/p} & \text{if } p \neq 0, \\ \left(\prod_{i=1}^{k}U_{ik}\right)^{1/k} & \text{if } p = 0, \end{cases}$$

and, with $\hat{\xi}_k^{H_0} \equiv \hat{\xi}_k^{H}$, defined in (1.3), the class of MO_p EVI-estimators,

$$\hat{\xi}_{k}^{H_{p}} = \hat{\xi}_{k}^{H_{p}}(\underline{\mathbf{X}}_{n}) \coloneqq \begin{cases} \left(1 - M_{p}^{-p}(k)\right)/p & \text{if } p < 1/\xi, \\ \ln M_{0}(k) = \hat{\xi}_{k}^{H} & \text{if } p = 0. \end{cases}$$
(1.4)

The class of MO_p EVI-estimators in (1.4) depends on the tuning parameter $p \in \mathbb{R}$, it is highly flexible, it is scale-invariant but it is not invariant for changes in location, a property enjoyed by the EVI itself. To make the MO_p EVI-estimators in (1.4) location-invariant, it is sensible to use the *peaks over a random threshold* (PORT) methodology, introduced in Araújo-Santos et al. (2006), and further computationally studied in Gomes et al. (2008a). The PORT methodology is based on the sample of excesses over the random threshold $X_{n_q:n}$, $0 \le q < 1$, $n_q := \lfloor nq \rfloor + 1$, where $\lfloor x \rfloor$ denotes the integer part of x, i.e. it is based on the sample of size $n^{(q)} = n - n_q$,

$$\underline{\mathbf{X}}_{n}^{(q)} := (X_{n:n} - X_{n_{q:n}}, \dots, X_{n_{q+1:n}} - X_{n_{q:n}}).$$
(1.5)

The PORT-MO_p EVI-estimators are thus estimators with the same functional form of the EVI-estimators in (1.4), but with the original sample $\underline{\mathbf{X}}_n$ replaced by the sample of excesses $\underline{\mathbf{X}}_n^{(q)}$, in (1.5). Consequently, the PORT-MO_p EVI-estimators are given by $\hat{\boldsymbol{\xi}}_{\nu}^{\mathbf{H}_p^{(q)}} := \hat{\boldsymbol{\xi}}_{\nu}^{\mathbf{H}_p}(\mathbf{X}_n^{(q)})$.

Remark 1.1. We can have q = 0 whenever $F(\cdot)$ has a finite left endpoint (the random level can then be the minimum). Note that the choice q = 0 is appealing in practice, but should be used with care. Such a random threshold can indeed lead to under-estimation and even inconsistency (see, for instance Gomes et al., 2008a). Generally, we can have 0 < q < 1 (the random level is then an empirical quantile).

As a measure of comparison, the recent and promising *Pareto probability weighted moments* (PPWM) (see Caeiro and Gomes, 2011; Caeiro et al., 2014) will be considered. The PPWM EVI-estimators are consistent for $\xi < 1$, depend on the statistics,

$$\hat{a}_0(k) := \frac{1}{k} \sum_{i=1}^k X_{n-i+1:n}, \qquad \hat{a}_1(k) := \frac{1}{k} \sum_{i=1}^k \frac{i-1}{k-1} X_{n-i+1:n},$$

and are defined by

$$\hat{\xi}_k^{\text{PPWM}} = 1 - \frac{\hat{a}_1(k)}{\hat{a}_0(k) - \hat{a}_1(k)}, \quad 1 \le k < n.$$
(1.6)

PORT-PPWM EVI-estimators (see Caeiro et al., 2016) will also be included in the comparative studies to be developed in this article. In Section 2 a few details on second-order frameworks in EVT, reduced-bias estimation and asymptotic behavior of the estimators will be provided. Section 3 is dedicated to the finite sample properties of the EVI-estimators under study as well as their PORT-versions, done through a large-scale simulation study. Section 4 is devoted to a few final comments on the advantages of PORT EVI-estimators and on possible choices of the vector of tuning parameters.

2. Second-order frameworks for heavy RTFs, reduced-bias and PORT EVI-estimation

Let us consider for the reciprocal *right tail quantile function* (RTQF) associated with *F*, the notation $U(t) := F^{\leftarrow}(1 - 1/t)$, with $F^{\leftarrow}(y) := \inf\{x : F(x) \ge y\}$. For heavy right-tails, it is usual to work under the validity of a first-order condition

Download English Version:

https://daneshyari.com/en/article/1151536

Download Persian Version:

https://daneshyari.com/article/1151536

Daneshyari.com