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Statistics and Probability Letters

journal homepage: www.elsevier.com/locate/stapro

A random process related to a random walk on upper triangular matrices over a finite field

Martin Hildebrand

Department of Mathematics and Statistics, University at Albany, State University of New York, Albany, NY 12222, United States

ABSTRACT

ARTICLE INFO

Article history: Received 27 August 2013 Accepted 1 March 2014 Available online 11 March 2014

Keywords: Random process Upper triangular matrices

1. Introduction

A couple of random walks on $U_n(\mathbb{F}_q)$, the group of n by n upper triangular matrices with entries in a finite field \mathbb{F}_q and $n \ge 2$, have been studied. One walk, denoted by \mathcal{W}_1 , involves at each step adding a times row i to row i - 1 where i is chosen uniformly from $\{2, \ldots, n\}$ and a is chosen uniformly from \mathbb{F}_q . Another walk, denoted by \mathcal{W}_2 , involves at each step adding ± 1 times row i to row i - 1 where i chosen uniformly from $\{2, \ldots, n\}$ and each step adding the start with the identity matrix, and each step is chosen independent of earlier steps.

This article considers a random process related to a random walk on n by n upper triangular

matrices over a finite field \mathbb{F}_q where q is an odd prime. The walk starts with the identity,

and at each step, i is selected at random from $\{2, \ldots, n\}$ and either row i or the negative of

row *i* is added to row i - 1. This article shows that, for a given *q*, it takes order n^2 steps for the last column to get close to uniformly distributed over all possibilities for that column.

In Peres and Sly (2013), they have shown that a lazy version of W_1 is close to uniformly distributed on $U_n(\mathbb{F}_q)$ after order $n^2 \log q$ steps. An important part of their argument involves the last column of the matrix. In Stong (1995), he has shown that after order $n^3q^2 \log q$ steps, W_2 is close to uniformly distributed on $U_n(\mathbb{F}_q)$.

In this paper, we consider the last column of the matrix in the random walk W_2 . We show that if q is a given odd prime, then order n^2 steps suffice to make the last column close to uniformly distributed on all possibilities for the last column. The proof involves representation theory of \mathbb{F}_q^{n-1} and the Upper Bound lemma of Diaconis and Shahshahani.

2. Notation and background

Let \mathbb{F}_q be a finite field with q elements in it where q is an odd prime number. Let $U_n(\mathbb{F}_q)$ be the group of n by n upper triangular matrices with 1's on the diagonal, 0's below the diagonal, and entries in \mathbb{F}_q above the diagonal.

If *P* is a probability on a finite set *S* and *U* is the uniform probability on *S* (i.e. U(s) = 1/|S| for all $s \in S$), then we define the variation distance of *P* from *U* by

$$||P - U|| = \frac{1}{2} \sum_{s \in S} |P(s) - U(s)|.$$

It can be shown that

 $||P - U|| = \max_{A \subset S} |P(A) - U(A)|$

where $P(A) = \sum_{s \in A} P(s)$ and the maximum is over all subsets A of S.

E-mail address: mhildebrand@albany.edu.

http://dx.doi.org/10.1016/j.spl.2014.03.001 0167-7152/© 2014 Elsevier B.V. All rights reserved.







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If the set *S* is the set of elements of a finite group *G*, then the Upper Bound lemma of Diaconis and Shahshahani (described in Diaconis, 1988) can be used to bound ||P - U||.

Lemma 1 (Upper Bound Lemma of Diaconis and Shahshahani).

$$||P - U||^2 \le \frac{1}{4} \sum_{\rho}^* d_{\rho} \operatorname{Tr}(\hat{P}(\rho)\hat{P}(\rho)^*)$$

where the sum is over all irreducible representations ρ on *G* up to equivalence, d_{ρ} is the degree of ρ , the Fourier transform $\hat{P}(\rho) = \sum_{s \in C} P(s)\rho(s)$, the irreducible representations are assumed to be unitary, and $\hat{P}(\rho)^*$ is the conjugate transpose of $\hat{P}(\rho)$.

3. Main result

We shall define random variables $X_0, X_1, X_2, ...$ on \mathbb{F}_q^{n-1} as follows. Let X_0 be the column vector with n - 1 0's. Let

Let X_0 be the column vector with n -

$$X_{m+1} = A_m X_m + b_m$$

where $(A_0, b_0), (A_1, b_1), ...$ are i.i.d. such that

$$P((A_0, b_0) = (M_{\ell}^+, c_{\ell})) = P((A_0, b_0) = (M_{\ell}^-, -c_{\ell})) = \frac{1}{2(n-1)}$$

for $\ell = 2, ..., n$, and M_{ℓ}^+ , M_{ℓ}^- , and c_{ℓ} are defined as follows. M_n^+ and M_n^- are (n-1) by (n-1) identity matrices. c_n is a column vector with n - 2 0's followed by a 1. M_{ℓ}^+ , for $2 \le \ell \le n - 1$, is (n - 1) by (n - 1) matrix with diagonal entries 1, with the entry in position $(\ell - 1, \ell)$ being 1, and all other entries 0. M_{ℓ}^- , for $2 \le \ell \le n - 1$, is (n - 1) by (n - 1) matrix with diagonal entries 1, with the entry in position $(\ell - 1, \ell)$ being -1, and with all other entries 0. c_{ℓ} , for $2 \le \ell \le n - 1$, is the zero vector with n - 1 entries being in one column.

Note that X_m is the last column of the upper triangular matrix after *m* steps of the random walk W_2 . Also note that if n = 2, this random process is a simple random walk on \mathbb{F}_q .

Let P_m be the probability on \mathbb{F}_q^{n-1} such that for all $s \in \mathbb{F}_q^{n-1}$, $P_m(s)$ is the probability that $X_m = s$. The main result is as follows.

Theorem 1. Suppose q is a given odd prime. Then for some constant $c_q > 0$, if $m > c_q n^2$, then $||P_m - U|| \to 0$ as $n \to \infty$.

We assume $n \ge 3$ in the proof.

To prove this theorem, we shall use a recurrence relation involving the Fourier transforms of P_m , and then we use this recurrence relation, some arguments involving discrete probability, and the Upper Bound Lemma.

4. Recurrence relation for Fourier transforms

First observe

$$P(X_{m+1}=j) = \frac{1}{2(n-1)} \sum_{\ell=2}^{n} (P(M_{\ell}^{+}X_{m} + c_{\ell} = j) + P(M_{\ell}^{-}X_{m} - c_{\ell} = j)).$$

For each $k \in \mathbb{F}_q^{n-1}$, let $\hat{P}_{m+1}(k)$ denote $\hat{P}_{m+1}(\rho_k)$ where $\rho_k(j) = \omega^{k^T j}$ where $\omega = e^{2\pi i/q}$ and k^T is the transpose of k. Then

$$\begin{split} \hat{P}_{m+1}(k) &= \sum_{j} P(X_{m+1} = j) \omega^{k^{l} j} \\ &= \sum_{j} \frac{1}{2(n-1)} \left(\sum_{\ell=2}^{n} (P(M_{\ell}^{+}X_{m} + c_{\ell} = j) + P(M_{\ell}^{-}X_{m} - c_{\ell} = j)) \right) \omega^{k^{T} j} \\ &= \sum_{\ell=2}^{n} \frac{1}{2(n-1)} \left(\sum_{j} P(X_{m} = (M_{\ell}^{+})^{-1}(j-c_{\ell})) \omega^{k^{T} j} + \sum_{j} P(X_{m} = (M_{\ell}^{-})^{-1}(j+c_{\ell})) \omega^{k^{T} j} \right). \end{split}$$

Note that

$$\sum_{j} P(X_{m} = (M_{\ell}^{+})^{-1}(j - c_{\ell}))\omega^{k^{T}j} = \sum_{j} P(X_{m} = (M_{\ell}^{+})^{-1}j)\omega^{k^{T}(j+c_{\ell})}$$
$$= \sum_{j} P(X_{m} = j)\omega^{k^{T}(M_{\ell}^{+})^{T}+c_{\ell}}$$
$$= \sum_{j} P(X_{m} = j)\omega^{((M_{\ell}^{+})^{T}k)^{T}j}\omega^{k^{T}c_{\ell}}$$
$$= \hat{P}_{m}((M_{\ell}^{+})^{T}k)\omega^{k^{T}c_{\ell}}.$$

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