



A remark on the optimal transport between two probability measures sharing the same copula



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ABSTRACT

We study the optimal transport between two probability measures on \mathbb{R}^n sharing the same copula C . We investigate the optimality of the image of the probability measure dC by the vectors of pseudo-inverses of marginal distributions.

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1. Optimal transport between two probability measures sharing the same copula

Given two probability measures μ and ρ , the optimal transport theory aims at minimizing $\int c(x, y)v(dx, dy)$ over all couplings v with first marginal $v \circ ((x, y) \mapsto x)^{-1} = \mu$ and second marginal $v \circ ((x, y) \mapsto y)^{-1} = \rho$ for a measurable non-negative cost function c . We use the notation $v \ll_{\rho}^{\mu}$ for such couplings. In the present note, we are interested in the particular case of the so-called Wasserstein distance between two probability measures μ and ρ on \mathbb{R}^n :

$$\mathcal{W}_{p,q}(\mu, \rho) = \inf_{v \ll_{\rho}^{\mu}} \left(\int_{\mathbb{R}^n \times \mathbb{R}^n} \|x - y\|_q^p v(dx, dy) \right)^{1/p} \tag{1.1}$$

obtained for the choice $c(x, y) = \|x - y\|_q^p$. Here \mathbb{R}^n is endowed with the norm $\|(x_1, \dots, x_n)\|_q = (\sum_{i=1}^n |x_i|^q)^{1/q}$ for $q \in [1, +\infty)$ whereas $p \in [1, +\infty)$ is the power of this norm in the cost function.

In dimension $n = 1$, $\|x\|_q = |x|$ so that the Wasserstein distance does not depend on q and is simply denoted by \mathcal{W}_p . Moreover, the optimal transport is given by the inversion of the cumulative distribution functions: whatever $p \in [1, +\infty)$, an optimal coupling is the image of the Lebesgue measure on $(0, 1)$ by $u \mapsto (F_{\mu}^{-1}(u), F_{\rho}^{-1}(u))$ where for $u \in (0, 1)$, $F_{\mu}^{-1}(u) = \inf\{x \in \mathbb{R} : \mu((-\infty, x]) \geq u\}$ and $F_{\rho}^{-1}(u) = \inf\{x \in \mathbb{R} : \rho((-\infty, x]) \geq u\}$ (see for instance Theorem 3.1.2 in Rachev and Rüschendorf (1998)). This implies that $\mathcal{W}_p^p(\mu, \rho) = \int_{(0,1)} |F_{\mu}^{-1}(u) - F_{\rho}^{-1}(u)|^p du$.

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In higher dimensions, according to Sklar’s theorem (see for instance Theorem 2.10.11 in Nelsen (2006)),

$$\mu \left(\prod_{i=1}^n (-\infty, x_i] \right) = C (\mu_1((-\infty, x_1]), \dots, \mu_n((-\infty, x_n]))$$

where we denote by $\mu_i = \mu \circ ((x_1, \dots, x_n) \mapsto x_i)^{-1}$ the i th marginal of μ and C is a copula function i.e. $C(u_1, \dots, u_n) = m \left(\prod_{i=1}^n [0, u_i] \right)$ for some probability measure m on $[0, 1]^n$ with all marginals equal to the Lebesgue measure on $[0, 1]$. The copula function C is uniquely determined on the product of the ranges of the marginal cumulative distribution functions $x_i \mapsto \mu_i((-\infty, x_i])$. In particular, when the marginals μ_i do not weight points, the copula C is uniquely determined. Sklar’s theorem shows that the dependence structure associated with μ is encoded in the copula function C . Last, we give the well-known Fréchet–Hoeffding bounds

$$\forall u_1, \dots, u_n \in [0, 1], \quad C_n^-(u_1, \dots, u_n) \leq C(u_1, \dots, u_n) \leq C_n^+(u_1, \dots, u_n)$$

that hold for any copula function C with $C_n^+(u_1, \dots, u_n) = \min(u_1, \dots, u_n)$ and $C_n^-(u_1, \dots, u_n) = (u_1 + \dots + u_n - n + 1)^+$ (see Nelsen (2006), Theorem 2.10.12 or Rachev and Rüschendorf (1998), section 3.6). We recall that the copula C_n^+ is the n -dimensional cumulative distribution function of the image of the Lebesgue measure on $[0, 1]$ by $\mathbb{R} \ni x \mapsto (x, \dots, x) \in \mathbb{R}^n$. Also the copula C_2^- is the two-dimensional cumulative distribution function of the image of the Lebesgue measure on $[0, 1]$ by $\mathbb{R} \ni x \mapsto (x, 1 - x) \in \mathbb{R}^2$ and, for $n \geq 3$, C_n^- is not a copula.

In dimension $n = 1$, the unique copula function is $C(u) = u$ and therefore the optimal coupling between μ and ρ , which necessarily share this copula, is the image of the probability measure dC by $u \mapsto (F_\mu^{-1}(u), F_\rho^{-1}(u))$. It is therefore natural to wonder whether, when μ and ρ share the same copula C in higher dimensions, the optimal coupling is still the image of the probability measure dC by $(u_1, \dots, u_n) \mapsto (F_{\mu_1}^{-1}(u_1), \dots, F_{\mu_n}^{-1}(u_n), F_{\rho_1}^{-1}(u_1), \dots, F_{\rho_n}^{-1}(u_n))$. We denote by $\mu \diamond \rho$ this probability law on \mathbb{R}^{2n} . It turns out that the picture is more complicated than in dimension one because of the choice of the index q of the norm: optimality is guaranteed only when $p = q$ i.e. when the cost $\|x - y\|^p$ in (1.1) may be decomposed as the sum of coordinatewise costs.

Proposition 1.1. *Let $n \geq 2$, μ and ρ be two probability measures on \mathbb{R}^n sharing the same copula C and $\mathcal{W}_{p,q}(\mu, \rho) = \inf_{\nu \ll \mu \otimes \rho} \left(\int_{\mathbb{R}^n \times \mathbb{R}^n} \|x - y\|_q^p d\nu(dx, dy) \right)^{1/p}$.*

- If $p = q$, then an optimal coupling between μ and ρ is given by $\nu = \mu \diamond \rho$ and

$$\mathcal{W}_{p,p}^p(\mu, \rho) = \int_{[0,1]^n} \sum_{i=1}^n |F_{\mu_i}^{-1}(u_i) - F_{\rho_i}^{-1}(u_i)|^p dC(u_1, \dots, u_n) = \int_{[0,1]^n} \sum_{i=1}^n |F_{\mu_i}^{-1}(u) - F_{\rho_i}^{-1}(u)|^p du.$$

- If $p \neq q$, the coupling $\mu \diamond \rho$ is in general no longer optimal. For $p < q$, if $C \neq C_n^+$, we can construct probability measures μ and ρ on \mathbb{R}^n admitting C as their unique copula such that

$$\left(\int_{\mathbb{R}^n \times \mathbb{R}^n} \|x - y\|_q^p \mu \diamond \rho(dx, dy) \right)^{1/p} > \mathcal{W}_{p,q}(\mu, \rho).$$

For $p > q$, the same conclusion holds if $n \geq 3$ or $n = 2$ and $C \neq C_2^-$.

Remark 1.2. Let μ and ρ be two probability measures on \mathbb{R}^n and $\nu \ll \mu \otimes \rho$. For $n = 1$, ν is said to be comonotonic if $\nu((-\infty, x], (-\infty, y]) = C_2^+(\mu((-\infty, x]), \rho((-\infty, y]))$. Puccetti and Scarsini (2010) investigate several extensions of this notion for $n \geq 2$. In particular, they say that ν is π -comonotonic (resp. c -comonotonic) if μ and ρ have a common copula and $\nu = \mu \diamond \rho$ (resp. ν maximizes $\int_{\mathbb{R}^n \times \mathbb{R}^n} \langle x, y \rangle \tilde{\nu}(dx, dy)$ over all the coupling measures $\tilde{\nu} \ll \mu \otimes \rho$). Looking at some connections between their different definitions of comonotonicity, they show in Lemma 4.4 that π -comonotonicity implies c -comonotonicity. Since

$$\int_{\mathbb{R}^n \times \mathbb{R}^n} \|x - y\|_2^2 \tilde{\nu}(dx, dy) = \int_{\mathbb{R}^n} \|x\|_2^2 \mu(dx) + \int_{\mathbb{R}^n} \|y\|_2^2 \rho(dy) - 2 \int_{\mathbb{R}^n \times \mathbb{R}^n} \langle x, y \rangle \tilde{\nu}(dx, dy),$$

this yields our result in the case $p = q = 2$.

2. Proof of Proposition 1.1

The optimality in the case $q = p$, follows by choosing $d_1 = \dots = d_n = d'_1 = \dots = d'_n = d''_1 = \dots = d''_n = 1$, $c_i(y_i, z_i) = |y_i - z_i|^p$, $\alpha = dC$, and $\varphi_i = F_{\mu_i}^{-1}$, $\psi_i = F_{\rho_i}^{-1}$ in the following Lemma.

Lemma 2.1. *Let α be a probability measure on $\mathbb{R}^{d_1} \times \mathbb{R}^{d_2} \times \dots \times \mathbb{R}^{d_n}$ with respective marginals $\alpha_1, \dots, \alpha_n$ on $\mathbb{R}^{d_1}, \dots, \mathbb{R}^{d_n}$ and $\varphi_i : \mathbb{R}^{d_i} \rightarrow \mathbb{R}^{d'_i}$, $\psi_i : \mathbb{R}^{d_i} \rightarrow \mathbb{R}^{d''_i}$ and $c_i : \mathbb{R}^{d'_i} \times \mathbb{R}^{d''_i} \rightarrow \mathbb{R}_+$ be measurable functions such that*

$$\forall i \in \{1, \dots, n\}, \quad \inf_{\substack{\alpha_i \circ \varphi_i^{-1} \\ \nu_i < \alpha_i \circ \psi_i^{-1}}} \int_{\mathbb{R}^{d'_i} \times \mathbb{R}^{d''_i}} c_i(y_i, z_i) \nu_i(dy_i, dz_i) = \int_{\mathbb{R}^{d_i}} c_i(\varphi_i(x_i), \psi_i(x_i)) \alpha_i(dx_i). \tag{2.1}$$

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