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# Threshold behaviors of a random constraint satisfaction problem with exact phase transitions $\stackrel{\scriptscriptstyle \times}{\scriptscriptstyle \propto}$

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#### 1. Introduction

Constraint satisfaction problems (CSPs) arise in the disciplines of theoretical computer science, discrete mathematics and statistical physics. Generally speaking, a CSP consists of n variables and m constraints. Each variable takes values from a nonempty domain D with size |D| = d. Each constraint is defined on a subset of k distinct variables and specifies the incompatible values for these variables. An assignment of the n variables satisfying the mconstraints simultaneously is a solution of the CSP. Given a CSP instance, the problem is to find such an assignment or to show that it is unsatisfiable. Accumulative empirical evidences suggest that CSPs undergo phase transitions at some critical values. The k-satisfiability problem (k-SAT) is perhaps the most famous example of CSPs [1].

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#### ABSTRACT

We consider a random constraint satisfaction problem named model RB, which exhibits a sharp satisfiability phase-transition phenomenon when the control parameters pass through the critical values denoted by  $r_{cr}$  and  $p_{cr}$ . Using finite-size scaling analysis, we bound the width of the transition region for finite problem size n, which might be the first rigorous study on the threshold behaviors of NP-complete problems.

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For k = 2, the problem is in P and it has been analytically shown that the threshold locates at  $\alpha = 1$  ( $\alpha \equiv m/n$ ) [2], which is further strengthened in [3] and [4]. To gain a better understanding of the changes of phase-transition region, finite-size scaling analysis, a method from statistical physics was introduced. For random 2-SAT, how the width of the transition region narrows with increasing problem size is shown in [5] and [6]. However, for random *k*-SAT ( $k \ge 3$ ), which is NP-complete, rigorous analysis appears to be difficult to locate the exact values of critical points. Loose upper and lower bounds have been obtained. Experimental evidence, however, strongly suggests a threshold occurs at  $\alpha \approx 4.2667$  [7] for random 3-SAT.

Model RB is a random CSP proposed by Xu and Li [8] to overcome the trivial unsatisfiability of standard CSP models. In model RB, first, we construct a constraint by randomly choosing k ( $k \ge 2$ ) distinct variables among n ones, and then uniformly select  $pd^k$  ( $d = n^\alpha$ ) incompatible values for the k variables. Next, we repeat this process to obtain  $m = rn \ln n$  independently chosen constraints and their associated sets of incompatible values. We take the logical AND of the m constraints, which gives an instance

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of model RB. The probability of a random instance I being satisfiable is denoted by Pr(I is SAT), it is shown in [8] that

**Theorem 1.1.** Let  $r_{cr} = -\frac{\alpha}{\ln(1-p)}$ . If  $\alpha > \frac{1}{k}$ ,  $0 are two constants and <math>k \ge \frac{1}{1-p}$ , then

 $\lim_{n \to \infty} \mathbf{Pr}(I \text{ is SAT}) = 1, \quad \text{if } r < r_{cr}$  $\lim_{n \to \infty} \mathbf{Pr}(I \text{ is SAT}) = 0, \quad \text{if } r > r_{cr}$ 

**Theorem 1.2.** Let  $p_{cr} = 1 - e^{-\frac{\alpha}{r}}$ . If  $\alpha > \frac{1}{k}$ , r > 0 are two constants and  $ke^{-\frac{\alpha}{r}} \ge 1$ , then

$$\lim_{n \to \infty} \mathbf{Pr}(l \text{ is SAT}) = 1, \quad \text{if } p < p_{cr}$$
$$\lim_{n \to \infty} \mathbf{Pr}(l \text{ is SAT}) = 0, \quad \text{if } p > p_{cr}$$

Therefore, the satisfiability phase transition of model RB is exactly established. Moreover, it has been shown both theoretically [9] and experimentally [10] that almost all instances of model RB are hard to solve at the thresholds. Since hard instances can be used as benchmarks for evaluating the performance of algorithms, the instances of model RB have been widely used in various algorithm competitions and many research papers in the past few years. So studies on the threshold behaviors of model RB can provide further understanding of the hardness of these instances at the thresholds.

In this paper, we use finite-size scaling analysis to characterize the threshold behaviors of model RB with finite problem size n, which give direct evidence for the broadening of the transition region due to finite-size effects.

#### 2. Main results

Suppose  $\delta$  is an arbitrarily small constant. We can obtain the following results under the same conditions of Theorem 1.1.

**Theorem 2.1.** Let  $r_u = r_{cr} + \Theta(\frac{1}{n \ln n})$ , then

 $\mathbf{Pr}(I \text{ is SAT}) \leqslant \delta, \quad \text{if } r \geqslant r_u.$ 

**Theorem 2.2.** Let 
$$r_l = r_{cr} - \Theta(\frac{1}{n^{1-\nu} \ln n})$$
, then

$$\mathbf{Pr}(I \text{ is SAT}) \ge 1 - \delta, \quad \text{if } r \le r_l$$

where the constants implicit in  $\Theta$  depend on  $\delta$  and  $0 < \nu < 1$  is a constant.

It is shown that the threshold of model RB scales as  $\frac{1}{n \ln n}$  and  $\frac{1}{n^{1-\nu} \ln n}$  from above and below respectively, which bounds the width of the transition region accordingly. Clearly  $\Theta(\frac{1}{n \ln n})$  and  $\Theta(\frac{1}{n^{1-\nu} \ln n})$  vanish as *n* approaches infinity. So the exact critical point  $r_{cr}$  can be identified by the fact that  $\delta$  is arbitrarily small. Similar results about the control parameter *p* can also be obtained by analogy with Theorems 2.1 and 2.2. In this paper, we mainly focus on the results about the control parameter *r*.

#### 3. Proof of the results

For a random instance *I*, let  $S(I) = \{\sigma : \sigma \text{ satisfies } I\} \subseteq D^n$  be the set of solutions of *I*, where  $\sigma$  is an assignment of the *n* variables. Let X = |S(I)| denote the number of solutions of *I*, we will prove Theorem 2.1 by Markov inequality, and use the second moment method to prove Theorem 2.2.

**Proof of Theorem 2.1.** We begin by calculating the expectation of *X* 

$$\mathbf{E}(X) = d^n \operatorname{\mathbf{Pr}}(\sigma \text{ satisfies } I) = n^{\alpha n} (1-p)^{rn \ln n}$$
(1)

Assume that  $\mathbf{E}(X) \leq \delta$ , we have

$$\left[\alpha + r\ln(1-p)\right]n\ln n \leqslant \ln\delta \tag{2}$$

Hence we can obtain

$$r \ge r_{cr} + \frac{\ln\delta}{n\ln n\ln(1-p)} \equiv r_{cr} + \Theta\left(\frac{1}{n\ln n}\right)$$
(3)

Therefore, using Markov inequality  $\mathbf{Pr}(I \text{ is SAT}) \leq \mathbf{E}(X)$ , we can immediately get that if  $r \geq r_u$ , then  $\mathbf{Pr}(I \text{ is SAT}) \leq \delta$ .  $\Box$ 

**Proof of Theorem 2.2.** Let  $Pr(\sigma_1, \sigma_2 \text{ satisfy } I)$  denote the probability that a pair of assignments  $\sigma_1$ ,  $\sigma_2$  satisfy I simultaneously. So we can have

$$\mathbf{E}(X^2) = \sum_{t=0}^{n} d^n \binom{n}{t} (d-1)^{n-t} \operatorname{\mathbf{Pr}}(\sigma_1, \sigma_2 \text{ satisfy } I)$$
(4)

where *t* denotes the number of variables for which  $\sigma_1$  and  $\sigma_2$  take the same values, and

**Pr**( $\sigma_1, \sigma_2$  satisfy *I*)

$$= \left[ (1-p)\frac{\binom{t}{k}}{\binom{n}{k}} + \frac{\binom{d^{k}-2}{pd^{k}}}{\binom{d^{k}}{pd^{k}}} \left(1 - \frac{\binom{t}{k}}{\binom{n}{k}}\right) \right]^{rn\ln n}$$
(5)

**Explanation of (5):** For each constraint in *I*, the probability of  $\sigma_1$  and  $\sigma_2$  being assigned the same values to the *k* variables is  $\binom{t}{k} / \binom{n}{k}$ , in this case, the probability of  $\sigma_1$  and  $\sigma_2$  satisfying the constraint is 1 - p; otherwise the probability is  $1 - \binom{t}{k} / \binom{n}{k}$ , and the satisfying probability is  $\binom{d^k-2}{pd^k} / \binom{d^k}{pd^k}$ . By the condition of  $\alpha > \frac{1}{k}$  in Theorem 1.1, we can rewrite (5) as

**Pr**( $\sigma_1, \sigma_2$  satisfy *I*)

$$= \left[ (1-p)A(t) + (1-p)^2 (1-A(t)) \right]^{rn\ln n}$$
$$\times \left( 1 + O\left(\frac{1}{n}\right) \right) \tag{6}$$

where  $A(t) = (\frac{t}{n})^k + \frac{g(t)}{n}$  with  $g(t) = \frac{k(k-1)}{2}(\frac{t}{n})^{k-1}(\frac{t}{n}-1)$ . Clearly  $A(t) \leq (\frac{t}{n})^k$  for  $g(t) \leq 0$  with  $t \in [0, n]$ . Consequently, we use (1), (4) and (6) to have Download English Version:

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