Original Article

# Quartile ranked set sampling for estimating the distribution function 

CrossMark

Amer Ibrahim Al-Omari

Al al-Bayt University, Faculty of Science, Department of Mathematics, Mafraq, Jordan

Received 24 January 2014; revised 5 July 2014; accepted 15 April 2015
Available online 21 May 2015

## Keywords

Distribution function;
Quartile ranked set sampling;
Simple random sampling;
Ranked set sampling;
Relative efficiency


#### Abstract

Quartile ranked set sampling (QRSS) method is suggested by Muttlak (2003) for estimating the population mean. In this article, the QRSS procedure is considered to estimate the distribution function of a random variable. The proposed QRSS estimator is compared with its counterparts based on simple random sampling (SRS) and ranked set sampling (RSS) schemes. It is found that the suggested estimator of the distribution function of a random variable $X$ for a given $x$ is biased and more efficient than its competitors using SRS and RSS.


2010 Mathematics Subject Classification: 62G30; 62D05; 62G05; 62G07
Copyright 2015, Egyptian Mathematical Society. Production and hosting by Elsevier B.V. This is an open access article under the CC BY-NC-ND license (http://creativecommons.org/licenses/by-nc-nd/4.0/).

## 1. Introduction

Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample of size $n$ having probability density function (pdf) $f(x)$ and cumulative distribution function (cdf) $F(x)$, with a finite mean $\mu$ and variance $\sigma^{2}$. Let $X_{11 i}, X_{12 i}, \ldots, X_{1 n i} ; X_{21 i}, X_{22 i}, \ldots, X_{2 n i} ; \ldots ; X_{n 1 i}, X_{n 2 i}, \ldots, X_{n n i}$ be $n$ independent simple random samples each of size $n$ in the $i$ th cycle $(i=1,2, \ldots, m)$.

Let $F_{S R S}(x)$ denote the empirical distribution function of a simple random sample $X_{1}, X_{2}, \ldots, X_{n m}$ from $F(x)$. Bahadur [2] showed that $F_{S R S}(x)$ has the following properties:

E-mail address: alomari_amer@yahoo.com Peer review under responsibility of Egyptian Mathematical Society.


1. $F_{S R S}(x)$ is an unbiased estimator of $F(x)$ for a given $x$.
2. $\operatorname{Var}\left[F_{S R S}(x)\right]=\frac{1}{m n} F(x)[1-F(x)]$.
3. $F_{S R S}(x)$ is a consistent estimator of $F(x)$.

The RSS was first suggested by McIntyre [3] as a method for estimating the mean of pasture and forage yields. The RSS is a useful method when the sampling units can be easily ranked than quantified. McIntyre proposed the ranked set sample mean as an estimator of the population mean and showed that the RSS mean estimator is unbiased and is more efficient than the SRS counterpart.

The RSS can be described as follows: randomly select $n$ sets each of size $n$ from the target population. Then, visually rank the units within each sample with respect to the variable of interest. From the first set of $n$ units the smallest ranked unit is selected. From the second set of $n$ units the second smallest ranked unit is selected. The process is continued until the largest ranked unit is measured from the $n$th set. To increase the

Table 1 The relative precision of $F_{Q R S S}(x)$ with respect to $F_{S R S}(x)$ and bias values of $F_{Q R S S}(x)$ for $4 \leq n \leq 11$.

| $F(x)$ |  | $n=4$ | $n=6$ | $n=8$ | $n=10$ | $n=5$ | $n=7$ | $n=9$ | $n=11$ |
| :--- | :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 0.01 | RP | 0.5270 | 7.9487 | 5.3923 | $\mathbf{9 . 1 6 9 0}$ | $\mathbf{1 1 . 0 2 4 8}$ | 7.0763 | 4.5962 | 8.2674 |
|  | Bias | 0.0090 | -0.0093 | -0.0088 | -0.0099 | -0.0096 | -0.0092 | -0.0084 | -0.0099 |
| 0.10 | RP | 0.6852 | $\mathbf{1 . 4 5 6 6}$ | 1.1587 | 1.2156 | $\mathbf{1 . 7 0 2 4}$ | 1.4095 | 1.1780 | 1.1780 |
|  | Bias | 0.0718 | -0.0429 | -0.0066 | -0.0650 | -0.0655 | -0.0356 | 0.0006 | -0.0592 |
| 0.20 | RP | 0.9939 | $\mathbf{1 . 3 2 6 0}$ | 1.0919 | 1.2896 | 1.3138 | $\mathbf{1 . 4 3 8 8}$ | 1.1553 | 1.3873 |
|  | Bias | 0.0973 | -0.0254 | 0.0482 | -0.0396 | -0.0802 | -0.0142 | 0.0527 | -0.0257 |
| 0.30 | RP | 1.6994 | 1.6332 | 1.5037 | $\mathbf{1 . 7 3 9 7}$ | 1.2961 | 1.9075 | 1.7570 | $\mathbf{1 . 9 3 2 5}$ |
|  | Bias | 0.0849 | -0.0044 | 0.0727 | 0.0098 | -0.0673 | 0.0062 | 0.0679 | 0.0201 |
| 0.40 | RP | 2.9902 | 2.2115 | $\mathbf{3 . 3 2 0 0}$ | 2.9599 | 1.3908 | 2.5254 | $\mathbf{3 . 3 4 8 0}$ | 3.1859 |
|  | Bias | 0.0475 | 0.0041 | 0.0505 | 0.0224 | -0.0367 | 0.0103 | 0.0447 | 0.0265 |
| 0.50 | RP | 4.3694 | 2.5640 | $\mathbf{7 . 3 5 7 3}$ | 4.8399 | 1.4534 | 2.8914 | $\mathbf{5} .5273$ | 4.8417 |
|  | Bias | 0.0005 | 0.0008 | 0.0000 | -0.0005 | 0.0003 | 0.0005 | 0.0004 | -0.0003 |
| 0.60 | RP | 3.0824 | 2.2121 | $\mathbf{3 . 2 7 3 6}$ | 2.9261 | 1.3843 | 2.5467 | $\mathbf{3 . 3 9 1 7}$ | 3.2440 |
|  | Bias | -0.0481 | -0.0040 | -0.0508 | -0.0216 | 0.0354 | -0.0105 | -0.0443 | -0.0256 |
| 0.70 | RP | 1.7065 | 1.6410 | 1.5109 | $\mathbf{1 . 7 5 2 1}$ | 1.2757 | 1.8497 | 1.7220 | $\mathbf{1 . 9 5 3 9}$ |
|  | Bias | -0.0848 | 0.0051 | -0.0733 | -0.0088 | 0.0662 | -0.0067 | -0.0686 | -0.0194 |
| 0.80 | RP | 1.0080 | $\mathbf{1 . 3 7 4 8}$ | 1.1220 | 1.2898 | 1.2997 | $\mathbf{1 . 4 2 2 9}$ | 1.1603 | 1.3658 |
|  | Bias | -0.0960 | 0.0277 | -0.0483 | 0.0391 | 0.0812 | 0.0136 | -0.0536 | 0.0256 |
| 0.90 | RP | 0.6628 | $\mathbf{1 . 4 4 5 5}$ | 1.1659 | 1.2054 | $\mathbf{1 . 6 7 5 6}$ | 1.4006 | 1.1633 | 1.1870 |
|  | Bias | -0.0736 | 0.0428 | 0.0063 | 0.0650 | 0.0659 | 0.0352 | 0.0004 | 0.0592 |
| 0.99 | RP | 0.5126 | 7.3353 | 4.9833 | $\mathbf{9 . 4 1 0 9}$ | $\mathbf{1 1 . 8 3 3 6}$ | 7.3257 | 4.4370 | 8.6399 |
|  | Bias | -0.0098 | 0.0092 | 0.0086 | 0.0099 | 0.0096 | 0.0093 | 0.0084 | 0.0099 |

sample size, the whole process can be repeated $m$ times to obtain a set of size $n m$ units.

Let $X_{j(1: n) i}, X_{j(2: n) i}, \ldots, X_{j(n: n) i}$ be the order statistics of the $j$ th sample $X_{j 1 i}, X_{j 2 i}, \ldots, X_{j n i}(j=1,2, \ldots, n)$ in the $i$ th cycle $(i=1,2, \ldots, m)$. Then, the measured units $X_{1(1: n) i}, X_{2(2: n) i}, \ldots, X_{n(n: n) i}$ are denoted to the RSS. David and Nagaraja [4] showed that the cdf and the pdf of the $j$ th order statistic $X_{(j: n)}$ are given by
$F_{(j: n)}(x)=\sum_{i=j}^{n}\binom{n}{i}[F(x)]^{i}[1-F(x)]^{n-i},-\infty<x<\infty$,
and
$f_{(j: n)}(x)=\frac{n!}{(j-1)!(n-j)!}[F(x)]^{j-1}[1-F(x)]^{n-j} f(x)$.
The mean and the variance of $X_{(j: n)}$ are given by $\mu_{(j: n)}=$ $\int_{-\infty}^{\infty} x f_{(j: n)}(x) d x$ and $\sigma_{(j: n)}^{2}=\int_{-\infty}^{\infty}\left(x-\mu_{(j: n)}\right)^{2} f_{(j: n)}(x) d x$, respectively. Takahasi and Wakimoto [5] independently introduced the same method of RSS with mathematical theory and showed that
$f(x)=\frac{1}{n} \sum_{j=1}^{n} f_{(j: n)}(x), \quad \mu=\frac{1}{n} \sum_{j=1}^{n} \mu_{(j: n)}$, and

$$
\sigma^{2}=\frac{1}{n} \sum_{j=1}^{n} \sigma_{(j: n)}^{2}+\frac{1}{n} \sum_{j=1}^{n}\left(\mu_{(j: n)}-\mu\right)^{2}
$$

For a fixed $x$, Stokes and Sager [6] suggested an estimator for $F(x)$ using RSS as
$F_{R S S}(x)=\frac{1}{m n} \sum_{i=1}^{m} \sum_{j=1}^{n} I\left(X_{j(j: n) i} \leq x\right)$,
where $I(\cdot)$ is an indicator function. They proved the following:

1. $F_{R S S}(x)$ is an unbiased estimator for $F(x)$.
2. $\operatorname{Var}\left[F_{R S S}(x)\right]=\frac{1}{m n^{2}} \sum_{j=1}^{n} F_{(j: n)}(x)\left[1-F_{(j: n)}(x)\right]$.
3. $\frac{F_{R S S}(x)-E\left[F_{R S S}(x)\right]}{\sqrt{\operatorname{Var}\left[F_{R S S}(x)\right]}}$ converges in distribution to the standard normal as $m \rightarrow \infty$ when $x$ and $n$ are fixed.

For more about estimation of the distribution function in ranked set sampling methods see Stokes and Sager [6], Samawi and Al-Sagheer [7], Kim and Kim [8], Al-Saleh and Samuh [9], and Ghosh and Tiwari [10].

The rest of this paper is organized as follows. In Section 2, we introduced the suggested estimation of the distribution function using QRSS method. The performance of the new estimator against its SRS and RSS counterparts is given in Section 3. Section 4, is devoted for some inferences about $F(x)$. In Section 5 , some concluding remarks are provided.

## 2. Estimation of $F(x)$ using QRSS

The quartile ranked set sampling procedure as suggested by Muttlak [1] can be summarized as follows. Randomly select $n$ samples each of size $n$ units from the target population and rank the units within each sample with respect to the variable of interest. If the sample size $n$ is even, select and measure from the first $n / 2$ samples the $Q_{1}(n+1)$ th smallest ranked unit of each sample, i.e., the first quartile, and from the second $n / 2$ samples the $Q_{3}(n+1)$ th smallest ranked unit of each sample, i.e., the third quartile. Note that, we always take the nearest integer of $Q_{1}(n+1)$ th and $Q_{3}(n+1)$ th where $Q_{1}=25 \%$, and $Q_{3}=75 \%$. If the sample size $n$ is odd, select and measure from the first ( $n-1$ )/2 samples the $Q_{1}(n+1)$ th smallest ranked unit of each sample and from the other $(n-1) / 2$ samples the $Q_{3}(n+1)$ th smallest ranked unit of each sample, and from one sample the median for that sample. The cycle can be repeated $m$ times if needed to get a sample of size $n m$ units.

If the sample size $n$ is even, in the $i$ th cycle $(i=1,2, \ldots, m)$, let $X_{j\left(Q_{1}(n+1): n\right) i}$ be the $\left(Q_{1}(n+1)\right)$ th smallest ranked unit of the $j$ th sample $\left(j=1,2, \ldots, \frac{n}{2}\right)$, and $X_{j\left(Q_{3}(n+1): n\right) i}$ be the $\left(Q_{3}(n+1)\right)$ th smallest ranked unit of the $j$ th sample

# https://daneshyari.com/en/article/483512 

Download Persian Version:

## https://daneshyari.com/article/483512

## Daneshyari.com

