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# The correct regularity condition and interpretation of asymmetry in EGARCH $^{\ast}$

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#### HIGHLIGHTS

• EGARCH is derived from a random coefficient complex nonlinear moving average stochastic process that leads to its specification.

• Asymmetry and leverage are discussed.

• The correct regularity condition for asymmetry in EGARCH is derived.

• The correct interpretation is provided.

#### ARTICLE INFO

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#### $A \hspace{0.1in} B \hspace{0.1in} S \hspace{0.1in} T \hspace{0.1in} R \hspace{0.1in} A \hspace{0.1in} C \hspace{0.1in} T$

In the class of univariate conditional volatility models, the three most popular are the generalized autoregressive conditional heteroskedasticity (GARCH) model of Engle (1982) and Bollerslev (1986), the GIR (or threshold GARCH) model of Glosten et al. (1992), and the exponential GARCH (or EGARCH) model of Nelson (1990, 1991). For purposes of deriving the mathematical regularity properties, including invertibility that relates the standardized residuals to the returns shocks, to determine the likelihood function for estimation, and the statistical conditions to establish asymptotic properties, it is essential to understand the stochastic properties underlying the three univariate models. The random coefficient autoregressive process was used to obtain GARCH by Tsay (1987), an extension of which was used by McAleer (2014) to obtain GJR. A random coefficient complex nonlinear moving average process was used by McAleer and Hafner (2014) to obtain EGARCH. These models can be used to capture asymmetry, which denotes the different impacts on conditional volatility of positive and negative shocks of equal magnitude, and possibly also leverage, which is the negative correlation between returns shocks and subsequent shocks to volatility (see Black 1976). McAleer (2014) showed that asymmetry was possible for GIR, but not leverage. McAleer and Hafner (2014) showed that leverage was not possible for EGARCH. Surprisingly, the condition for asymmetry in EGARCH seem to have been ignored in the literature, or has concentrated on the incorrect parametric condition, with no clear explanation, and hence with associated unclear and misleading interpretations. The purpose of the paper is to derive the regularity condition for asymmetry in EGARCH, and to provide the correct interpretation. It is shown that, in practice, EGARCH always displays asymmetry, though not leverage.

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#### 1. Introduction

In the class of univariate conditional volatility models, the three most popular are the generalized autoregressive conditional







heteroskedasticity (GARCH) model of Engle (1982) and Bollerslev (1986), the GJR (or threshold GARCH) model of Glosten et al. (1992), and the exponential GARCH (or EGARCH) model of Nelson (1990, 1991). Each of these models is widely read, used in practice, and highly cited.

For purposes of deriving the mathematical regularity properties, including invertibility, to determine the likelihood function for estimation, and the statistical conditions to establish asymptotic properties, it is convenient to understand the stochastic properties underlying the three univariate models. The random coefficient autoregressive process was used to obtain GARCH by Tsay (1987), an extension of which was used by McAleer (2014) to obtain GJR. A random coefficient complex nonlinear moving average process was used by McAleer and Hafner (2014) to obtain EGARCH.

These models can be used to capture asymmetry, which denotes the different impacts on conditional volatility of positive and negative shocks of equal magnitude, and possibly also leverage, which is the negative correlation between returns shocks and subsequent shocks to volatility (see Black, 1976). McAleer (2014) showed that asymmetry was possible for GJR, but not leverage. McAleer and Hafner showed that leverage was not possible for EGARCH.

Surprisingly, the condition for asymmetry in EGARCH seems to have been ignored in the literature, or has concentrated on an incorrect condition, with no clear explanation, and hence with associated misleading interpretations. The purpose of the paper is to derive the regularity condition for asymmetry in EGARCH to provide the correct interpretation. It is shown that, in practice, EGARCH always displays asymmetry, though not leverage.

The paper is organized as follows. In Section 2, the GARCH and EGARCH models are derived from different underlying stochastic processes that lead to their derivation, the first from a random coefficient autoregressive process, and the second from a random coefficient complex nonlinear moving average process. Asymmetry and leverage are discussed in Section 3. The correct regularity condition for asymmetry in EGARCH is derived, and the correct interpretation is given, in Section 4. Some concluding comments are presented in Section 5.

#### 2. Stochastic processes underlying GARCH and EGARCH

#### 2.1. Random coefficient autoregressive process-GARCH

Consider the conditional mean of financial returns, as follows:

$$y_t = E\left(y_t | I_{t-1}\right) + \varepsilon_t,\tag{1}$$

where the financial returns,  $y_t = \Delta \log P_t$ , represent the logdifference in financial commodity prices,  $P_t$ ,  $I_{t-1}$  is the information set at time t - 1, and  $\varepsilon_t$  is a conditionally heteroskedastic error term, or returns shock. In order to derive conditional volatility specifications, it is necessary to specify the stochastic processes underlying the returns shocks,  $\varepsilon_t$ .

Now consider the random coefficient AR(1) process underlying the return shocks,  $\varepsilon_t$ :

$$\varepsilon_t = \phi_t \varepsilon_{t-1} + \eta_t \tag{2}$$

where

 $\phi_t \sim iid(0, \alpha), \alpha \ge 0$ ,

 $\eta_t \sim iid(0,\omega), \omega \ge 0,$ 

 $\eta_t = \varepsilon_t / \sqrt{h_t}$  is the standardized residual, with  $h_t$  defined below.

Tsay (1987) derived the ARCH(1) model of Engle (1982) from Eq. (2) as:

$$h_t \equiv E\left(\varepsilon_t^2 | I_{t-1}\right) = \omega + \alpha \varepsilon_{t-1}^2 \tag{3}$$

where  $h_t$  represents conditional volatility, and  $I_{t-1}$  is the information set available at time t - 1. A lagged dependent variable,  $h_{t-1}$ , is typically added to Eq. (3) to improve the sample fit:

$$h_t \equiv E\left(\varepsilon_t^2 | I_{t-1}\right) = \omega + \alpha \varepsilon_{t-1}^2 + \beta h_{t-1}.$$
(4)

From the specification of Eq. (2), it is clear that both  $\omega$  and  $\alpha$  should be positive as they are the unconditional variances of two different stochastic processes. Moreover, as GARCH is symmetric, there is no asymmetry or leverage.

Given the non-normality of the returns shocks, the Quasi-Maximum Likelihood Estimators (QMLE) of the parameters have been shown to be consistent and asymptotically normal in several papers. For example, Ling and McAleer (2003) showed that the QMLE for a generalized ARCH(p, q) (or GARCH(p, q)) is consistent if the second moment is finite. A sufficient condition for the QMLE of GARCH(1, 1) in Eq. (4) to be consistent and asymptotically normal is  $\alpha + \beta < 1$ .

#### 2.2. Random coefficient complex nonlinear moving average process-EGARCH

A conditional volatility model that can accommodate asymmetry is the EGARCH model of Nelson (1990, 1991). McAleer and Hafner (2014) showed that EGARCH, specifically EARCH(1) = EGARCH(1, 0), could be derived from a random coefficient complex nonlinear moving average (RCCNMA) process, as follows:

$$\varepsilon_t = \phi_t \sqrt{|\eta_{t-1}|} + \psi_t \sqrt{\eta_{t-1}} + \eta_t \tag{5}$$

where

 $\phi_t \sim iid (0, \alpha),$   $\psi_t \sim iid (0, \gamma),$   $\eta_t \sim iid (0, \varphi),$   $\sqrt{\eta_{t-1}}$  is a complex-valued function of  $\eta_{t-1},$ and  $\eta_t = \varepsilon_t / \sqrt{h_t}$  is the standardized residual. McAleer and Hafner (2014) show that the conditional variance of the squared returns shocks in Eq. (5) is:

$$h_t = E\left(\varepsilon_t^2 | I_{t-1}\right) = \omega + \alpha |\eta_{t-1}| + \gamma \eta_{t-1},\tag{6}$$

where it is clear from the RCCNMA process in Eq. (5) that all three parameters should be positive as they are the variances of three different stochastic processes. The constant  $\omega$  is not equivalent to the unconditional variance of  $\eta_t$ . In Eq. (6), it is assumed that  $h_t > 0$  as  $\eta_{t-1}$  could be negative, which is equivalent to assuming  $\alpha > \gamma$ . This is not a problem when the logarithmic approximation is used in Eq. (7).

Although the transformation of  $h_t$  in Eq. (6) is not logarithmic, the approximation given by:

 $\log h_t = \log (1 + (h_t - 1)) \approx h_t - 1$ 

can be used to replace  $h_t$  in Eq. (6) with  $1 + \log h_t$  to lead to EARCH(1) = EGARCH(1, 0):

$$\log h_{t} = E\left(\varepsilon_{t}^{2}|I_{t-1}\right) = (\omega - 1) + \alpha |\eta_{t-1}| + \gamma \eta_{t-1},$$
(7)

The use of an infinite lag for the RCCNMA process in Eq. (5) would yield the standard EGARCH model with lagged conditional volatility, namely EGARCH(1, 1):

$$\log h_{t} = E\left(\varepsilon_{t}^{2}|I_{t-1}\right) = (\omega - 1) + \alpha |\eta_{t-1}| + \gamma \eta_{t-1} + \beta \log h_{t-1}.$$
(8)

As EGARCH can be derived from a random coefficient complex nonlinear moving average (RCCNMA) process, it follows from the specification in Eq. (5) that there is no invertibility condition to transform the returns shocks,  $\varepsilon_t$ , to the standardized residuals,  $\eta_t$ . Download English Version:

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