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## A simple ordered data estimator for inverse density weighted expectations

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## Abstract

We consider estimation of means of functions that are scaled by an unknown density, or equivalently, integrals of conditional expectations. The "ordered data" estimator we provide is root  $n$ consistent, asymptotically normal, and is numerically extremely simple, involving little more than ordering the data and summing the results. No sample-size-dependent smoothing is required. A similarly simple estimator is provided for the limiting variance. The proofs include new limiting distribution results for functions of nearest-neighbor spacings. Potential applications include endogenous binary choice, willingness to pay, selection, and treatment models.  $\odot$  2005 Elsevier B.V. All rights reserved.

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## 1. Introduction

Given an iid dataset  $(x_1, y_1), \ldots, (x_n, y_n)$ , on a bivariate random variable  $(x, y)$ , we propose to estimate an expectation of the inverse density weighted form

$$
\theta = \mathcal{E}\left[\frac{y}{f(x)}\right],\tag{1}
$$

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where  $f(x)$  is the unknown density of the continuously distributed x. The "ordered data" estimator we provide possesses the rather surprising property of achieving root  $n$ consistency and asymptotic normality without requiring sample-size-dependent smoothing. It also offers the advantage of being numerically extremely simple, requiring little more than ordering the data and summing the results, thereby avoiding issues regarding the selection of smoothers such as bandwidths, kernels, etc. A similarly simple estimator is provided for the limiting variance.

Inverse density weighted estimation applies generically to the estimation of definite integrals of conditional expectations. Suppose

$$
\theta = \int_{x \in \mathcal{X}} \mathbf{E}(w|x) \, \mathrm{d}x \tag{2}
$$

for some random variable w, where  $\mathscr{X} \subset \text{supp}(x)$ . Then  $\theta$  can be rewritten as  $\theta = \text{E}[y/f(x)]$ with  $y = wI(x \in \mathcal{X})$ , and where I is the indicator function that equals one if its argument is true and zero otherwise.

A number of existing semiparametric estimators make use of inverse density weighted expectations either directly or indirectly, via their relationship with integrated conditional expectations. Examples include density weighted least squares ([Newey and Ruud, 1994](#page--1-0)), average derivative estimation (Härdle and Stoker, 1989), estimators of willingness-to-pay models and general estimators of moments from binomial data ([Lewbel, 1997; McFadden,](#page--1-0) [1999; Lewbel et al., 2005\)](#page--1-0), semiparametric estimators of consumer surplus [\(Hausman and](#page--1-0) [Newey, 1995; Newey, 1997\)](#page--1-0), some discrete choice, sample selection, and other latent variable model estimators [\(Lewbel, 1998, 2000, 2002\)](#page--1-0), entropy measures of dependence ([Hong and White, 2005](#page--1-0)) and semiparametric functional tests [\(Hall and Yatchew, 2005\)](#page--1-0).

Let  $(y_{[i]}, x_{[i]})$  denote the *i*th observation when the data are sorted in increasing order of x, so  $x_{ij}$  is the *i*th order statistic and  $y_{ij}$  is the concomitant statistic to  $x_{ij}$ . Let  $(y_i, x_i)$  denote the *i*th observation when the data are left unsorted. Let  $F(x)$  denote the unknown distribution function of x. We show that the numerically trivial "ordered data" estimator

$$
\hat{\theta} = \sum_{i=1}^{n-1} (y_{[i+1]} + y_{[i]}) (x_{[i+1]} - x_{[i]}) / 2
$$
\n(3)

is root *n* consistent and asymptotically normal. Specifically,  $n^{1/2}(\hat{\theta} - \theta) \stackrel{d}{\rightarrow} N(0, 3\sigma^2/2)$ , where

$$
\sigma^2 = \mathcal{E}\left[\frac{\text{Var}(y|x)}{f^2(x)}\right]
$$
(4)

and a consistent estimator of  $\sigma^2$  is the simple expression

$$
\hat{\sigma}^2 = \frac{n}{4} \sum_{i=1}^{n-1} (y_{[i+1]} - y_{[i]})^2 (x_{[i+1]} - x_{[i]})^2.
$$
\n(5)

For some intuition for this estimator, let x be a point that lies between  $x_{[i]}$  and  $x_{[i+1]}$  for some *i*, and let *y* be a corresponding point that lies between  $y_{[i]}$  and  $y_{[i+1]}$ . Then  $y \approx$  $(y_{[i+1]} + y_{[i]})/2$  and

$$
\frac{1}{f(x)} = \left(\frac{dF(x)}{dx}\right)^{-1} \approx \frac{x_{[i+1]} - x_{[i]}}{F(x_{[i+1]}) - F(x_{[i]})} \approx \frac{x_{[i+1]} - x_{[i]}}{1/n},\tag{6}
$$

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