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ℓ_p -norm based iterative adaptive approach for robust spectral analysis

Yuan Chen*, H.C. So, Weize Sun

Department of Electronic Engineering, City University of Hong Kong, Kowloon, Hong Kong, China

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ABSTRACT

Recently, the iterative adaptive approach (IAA) has been shown to be an effective nonparametric methodology for high-resolution spectral analysis. Its main idea is to reformulate the nonlinear frequency estimation problem as a linear model whose parameters are updated iteratively according to weighted least squares. Since the derivation of the IAA is based on ℓ_2 -norm, it cannot work properly in heavy-tailed noise environment. In this paper, a generalized version of IAA is derived to provide accurate spectral estimation in the presence of impulsive noise, which replaces the ℓ_2 -norm by the ℓ_p -norm where 1 . Simulation results are included to demonstrate the outlier resistance performance of the proposed algorithm.

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1. Introduction

Spectral analysis has been an important topic in science and engineering because many real-world signals are well described by the sinusoidal model. Basically, the frequency components of the observed data can be obtained by means of either parametric or nonparametric techniques [1]. In the parametric approach, the signal is assumed to satisfy a generating model with known functional form, which allows the derivation of the optimal spectral estimators. However, the performance of these methods degrades when there is a mismatch between the assumed and actual signal models. On the other hand, no assumptions are made about the data in the nonparametric approach. A conventional representative is the periodogram, which is based on the Fourier transform, but its resolution is fundamentally limited by the available observation length. Recently, the iterative adaptive approach (IAA) [2,3] provides a breakthrough in the nonparametric methodology because of its very high accuracy and

* Corresponding author. Tel.: +852 34422607.

E-mail address: qchenyuan00@126.com (Y. Chen).

0165-1684/\$-see front matter © 2013 Elsevier B.V. All rights reserved. http://dx.doi.org/10.1016/j.sigpro.2013.06.020 resolution. This method can be interpreted as reformulating the nonlinear spectral estimation problem as a linear model whose coefficients, representing amplitudes at different frequencies on a fine grid, are updated iteratively according to weighted least squares (WLS). Since its development is based on ℓ_2 -norm optimization, the IAA is not robust to heavy-tailed noise, which occurs in many fields such as wireless communications, radar and sonar [4]. Typical models for impulsive noise include α -stable noise [5], Gaussian mixture model (GMM) [4], and generalized Gaussian distribution (GGD) [6].

Existing methods for robust spectral analysis include [7–11]. With the use of Huber's minimax robust statistics [12], Katkovnik [7] has developed a maximum-likelihood (ML) type *M*-periodogram for spectral analysis in heavy-tailed noise. Its extension to time–frequency analysis for nonstationary signals is provided in [8]. Linear combination of order statistics, which is also called *L*-estimation, has been studied in [9] for robust transforms and time–frequency representations. In [11], it is proved that the least absolute deviation (LAD) estimator provides the ML performance in the presence of Laplacian white noise. Recently, the concept of LAD is generalized to ℓ_p -norm minimization where $1 \le p < 2$ in [11]. In this paper, by







utilizing ℓ_p -norm minimization and IAA, we derive an outlier-resistant version of IAA, which is referred to as ℓ_p -IAA, for impulsive noise environment.

The rest of this paper is organized as follows. The development of ℓ_p –IAA is presented in Section 2. We can see that the ℓ_p –IAA generalizes [2] as it reduces to the original IAA by setting p=2. Computer simulations in Section 3 show the effectiveness of the proposed scheme as well as its superiority over the IAA in the presence of both α –stable and GMM noises. Finally, conclusions are drawn in Section 4.

2. Proposed method

Without loss of generality, the observed signal sequence is modeled as

$$y_n = \sum_{l=1}^{L} \gamma_l e^{j\omega_l n} + q_n, \quad n = 0, ..., N-1$$
(1)

where γ_l and $\varpi_l \in (0,2\pi)$ are the complex amplitude and frequency of the *l*th tone, respectively, *L* is the number of sinusoids, and q_n is an isotropic independent and identically distributed (IID) heavy-tailed noise. It is assumed that the phases of $\{\gamma_l\}$ are independently and uniformly distributed in $[0, 2\pi]$ [1] and *L* is unknown. Our task is to find $\{\varpi_l\}$ from $\{y_n\}$.

Although frequency estimation corresponds to a nonlinear model, (1) can be reformulated as the following linear model [2]:

$$\mathbf{y} = \mathbf{A}\mathbf{x} \tag{2}$$

where $\mathbf{y} = [\mathbf{y}_0 \cdots \mathbf{y}_{N-1}]^T$ with ^{*T*} denotes transpose, $\mathbf{A} = [\mathbf{a}(\omega_0) \cdots \mathbf{a}(\omega_{K-1})]$ with $\mathbf{a}(\omega_k) = [a_0(\omega_k) \cdots a_{N-1}(\omega_k)]^T$, $a_n(\omega_k) = e^{jn\omega_k}$, $\omega_k = 2\pi k/K$, k = 0, ..., K-1, and $\mathbf{x} = [x_0 \cdots x_{K-1}]^T$. Note that the noise components $\{q_n\}$ are absorbed in \mathbf{x} . Here, the admissible frequency interval $[0,2\pi]$ is divided into *K* uniform grid points $\{\omega_k\}$ while x_k and $\mathbf{a}(\omega_k)$ are the amplitude and frequency-vector associated with ω_k , respectively. Assuming that *K* is chosen sufficiently large and in the absence of noise, we have

$$x_k = \begin{cases} \gamma_l, & \omega_k = \varpi_l, \quad l = 1, ..., L \\ 0 & \text{otherwise.} \end{cases}$$
(3)

To achieve outlier resistance, the proposed ℓ_p -IAA replaces the ℓ_2 -norm in [2] by ℓ_p -norm where $1 , and the conceptual estimate of <math>x_k$, denoted by \hat{x}_k , is

$$\hat{x}_k = \arg\min_{x_k} \|\mathbf{y} - \mathbf{a}(\omega_k) x_k\|_{\mathbf{Q}_k^{-1}}^p \tag{4}$$

where

$$\begin{aligned} \|\mathbf{y} - \mathbf{a}(\omega_{k}) x_{k}\|_{\mathbf{Q}_{k}^{-1}}^{p} &= \sum_{i=0}^{N-1} \sum_{j=0}^{N-1} [\mathbf{Q}_{k}^{-1}]_{i,j} (y_{i} - a_{i}(\omega_{k}) x_{k})^{*} \\ |y_{j} - a_{j}(\omega_{k}) x_{k}|^{p-2} (y_{j} - a_{j}(\omega_{k}) x_{k}) \\ &= (\mathbf{y} - \mathbf{a}(\omega_{k}) x_{k})^{H} \mathbf{Q}_{k}^{-1} \mathbf{W}_{k} (\mathbf{y} - \mathbf{a}(\omega_{k}) x_{k}). \end{aligned}$$
(5)

Here, ⁻¹, * and ^{*H*} denote matrix inverse, conjugate, and conjugate transpose, respectively, \mathbf{Q}_k^{-1} is the weighting matrix with (i, j) entry $[\mathbf{Q}_k^{-1}]_{i,j}$, and $\mathbf{W}_k = \text{diag}(|y_0 - a_0(\omega_k)x_k|^{p-2} \cdots |y_{N-1} - a_{N-1}(\omega_k)x_k|^{p-2})$ is a diagonal matrix. The \mathbf{Q}_k has

$$\mathbf{Q}_{k} = E\{(\mathbf{y} - \mathbf{a}(\omega_{k})\mathbf{x}_{k})(\mathbf{y} - \mathbf{a}(\omega_{k})\mathbf{x}_{k})^{H}\mathbf{W}_{k}\}$$
(6)

where *E* denotes the expectation operator.

Based on iteratively reweighted least squares, we solve (4) in an iterative manner and the estimate at the (l+1)th iteration is computed as

$$\hat{x}_{k}^{(l+1)} = \arg \min_{x_{k}} \{ (\mathbf{y} - \mathbf{a}(\omega_{k})x_{k})^{H} (\mathbf{Q}_{k}^{(l)})^{-1} \mathbf{W}_{k}^{(l)} (\mathbf{y} - \mathbf{a}(\omega_{k})x_{k}) \}$$

$$= \frac{\mathbf{a}^{H}(\omega_{k}) (\mathbf{Q}_{k}^{(l)})^{-1} \mathbf{W}_{k}^{(l)} \mathbf{y}}{\mathbf{a}^{H}(\omega_{k}) (\mathbf{Q}_{k}^{(l)})^{-1} \mathbf{W}_{k}^{(l)} \mathbf{a}(\omega_{k})}, \quad k = 0, ..., K-1$$
(7)

where $(\mathbf{Q}_k^{(l)})^{-1}\mathbf{W}_k^{(l)}$ is the reweighted matrix which is characterized by $\{\hat{x}_k^{(l)}\}$. In the Appendix, we have shown that (7) can be further simplified as

$$\hat{x}_{k}^{(l+1)} = \frac{\mathbf{a}^{H}(\omega_{k})(\mathbf{R}^{(l)})^{-1}\mathbf{W}_{k}^{(l)}\mathbf{y}}{\mathbf{a}^{H}(\omega_{k})(\mathbf{R}^{(l)})^{-1}\mathbf{W}_{k}^{(l)}\mathbf{a}(\omega_{k})}, \quad k = 0, ..., K-1$$
(8)

where

$$\mathbf{R}^{(l)} = \mathbf{A} \operatorname{diag}(|\hat{\mathbf{x}}_{0}^{(l)}|^{p} \cdots |\hat{\mathbf{x}}_{K-1}^{(l)}|^{p}) \mathbf{A}^{H}$$
(9)

is the robust covariance matrix constructed from $\{\hat{x}_k^{(l)}\}$ and

$$\mathbf{W}_{k}^{(l)} = \operatorname{diag}(|y_{0} - a_{0}(\omega_{k})\hat{x}_{k}^{(l)}|^{p-2} \cdots |y_{N-1} - a_{N-1}(\omega_{k})\hat{x}_{k}^{(l)}|^{p-2}).$$
(10)

The steps of the proposed algorithm are summarized in Table 1. It is worth noting that (8) reduces to the standard IAA when p=2. That is to say, the ℓ_p -IAA is a generalized version of [2].

The computational complexities of the proposed method as well as standard IAA and its fast implementation [14,15] are now examined. At each iteration, the numbers of multiplications required are $2KN^2 + KN + 2N^3$, $2KN^2 + KN + N^3$ and $N^2 + 24N \log_2(2N) + 3K \log_2(K)$, respectively. The additional computational cost of N^3 in ℓ_p –IAA over the conventional scheme is due to the multiplication of ($\mathbf{R}^{(l)}$)⁻¹ and $\mathbf{W}_k^{(l)}$. That is to say, when the number of iterations is kept identical in all schemes, the fast implementation is the most computationally efficient. Nevertheless, we can follow [14,15] to produce a fast implementation for the ℓ_p –IAA.

3. Simulation results

To evaluate the spectral estimation performance of ℓ_p -IAA, computer simulations have been conducted. All results are based on 100 independent runs and p=1.2 is selected. First, we examine the mean square frequency error (MSFE) performance of the proposed algorithm for a single tone. We also include the results of the standard IAA and Cramér-Rao lower bound (CRLB). The stopping criterion in the ℓ_p -IAA and IAA is chosen according to [2]. In the first test, the signal is generated according to (1) where

Table 1

Summary of proposed algorithm.

- (i) Initialize the values of $\{\hat{x}_k^{(0)}\}$ using discrete Fourier transform
- (ii) Compute $\mathbf{R}^{(l)}$ and $\mathbf{W}_{k}^{(l)}$ using (9) and (10) for k = 0, ..., K-1
- (iii) Update \hat{x}_k using (8) for k = 0, ..., K-1
- (iv) Repeat Steps (ii)-(iii) until a stopping criterion is reached

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