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Twisted trees and inconsistency of tree estimation when gaps are treated as missing data – The impact of model mis-specification in distance corrections [☆]

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ABSTRACT

Statistically consistent estimation of phylogenetic trees or gene trees is possible if pairwise sequence similarities can be converted to a set of distances that are proportional to the true evolutionary distances. Susko et al. (2004) reported some strikingly broad results about the forms of inconsistency in tree estimation that can arise if corrected distances are not proportional to the true distances. They showed that if the corrected distance is a concave function of the true distance, then inconsistency due to long branch attraction will occur. If these functions are convex, then two “long branch repulsion” trees will be preferred over the true tree – though these two incorrect trees are expected to be tied as the preferred true. Here we extend their results, and demonstrate the existence of a tree shape (which we refer to as a “twisted Farris-zone” tree) for which a single incorrect tree topology will be guaranteed to be preferred if the corrected distance function is convex. We also report that the standard practice of treating gaps in sequence alignments as missing data is sufficient to produce non-linear corrected distance functions if the substitution process is not independent of the insertion/deletion process. Taken together, these results imply inconsistent tree inference under mild conditions. For example, if some positions in a sequence are constrained to be free of substitutions and insertion/deletion events while the remaining sites evolve with independent substitutions and insertion/deletion events, then the distances obtained by treating gaps as missing data can support an incorrect tree topology even given an unlimited amount of data.

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1. Introduction

Distance-based methods are fast and statistically consistent estimators of tree topology if the input distances converge (with increasing data) to values that are proportional to the evolutionary distance between tips. An evolutionary distance is the number of substitution events that have occurred along the path separating two tips. Typically a distance correction procedure is applied to the observed sequence differences to obtain a more accurate estimate of the evolutionary distance between pairs of sequences. However, in many cases it is not possible to correctly account for the evolutionary processes which generated the data. In other words, it is not always possible to accurately estimate the evolutionary distance for pairwise measurements of dissimilarity.

In a pioneering paper, Susko et al. (2004) showed how model misspecification can lead to transformed evolutionary distances that are non-linear with respect to evolutionary distance (i.e. concave or convex), and for which there are regions of tree space for which neighbor joining will be inconsistent. We extend this result further (Theorem 1 in Appendix A) by showing how virtually all misspecified correction functions lead to (strong) inconsistency (an incorrect tree will be unambiguously favored by neighbor-joining). A main focus of this paper involves a particular study of model misspecification in distance corrections that treats gaps as missing data.

2. Model

For variants of the simplest model of sequence evolution (Jukes and Cantor, 1969), all nucleotides are equally exchangeable and the simple proportion of sites that differ, the *p*-distance, is a sufficient statistic for estimating an evolutionary distance. Under such a

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80 model, M_g , the expected p -distance between a pair of taxa is a
 81 function of the evolutionary distance (path length in the tree) t
 82 between the taxa, that is, we have $\mathbb{E}_g[p] = g(t)$, where the function
 83 g is a monotonically (strictly) increasing function of t which is ana-
 84 lytic (i.e. has a power series expansion, and so derivatives exist of
 85 all orders) and satisfies $g(0) = 0$. For example, for the Jukes–Cantor
 86 model we have $g(t) = \frac{3}{4}(1 - e^{-\frac{4}{3}t})$. If the distances are corrected
 87 under a (possibly different), fully exchangeable model, M_f , then
 88 the estimated evolutionary distance \hat{t} is usually computed from
 89 the p -distance by using the ‘plug-in’ formula $\hat{t} = f^{-1}(p)$.

90 Thus, for any generating model for which p converges in prob-
 91 ability towards its expected value $\mathbb{E}_g[p] = g(t)$ (e.g. i.i.d. site substi-
 92 tution models) the estimated evolutionary distance \hat{t} will converge
 93 towards $\bar{t} = h(t)$, where $h(t) = f^{-1}(g(t))$. Note here that both p and \hat{t}
 94 are random variables, while \bar{t} is simply a function of t . Notice that
 95 this ‘transformed’ evolutionary distance \bar{t} is not exactly the
 96 expected value of \hat{t} , even when $f = g$ (Tajima, 1993), since the
 97 expectation of a non-linear function of random variable is not gener-
 98 ally equal to the function evaluated at the expected value of that
 99 variable. Nevertheless, for any i.i.d. site substitution model, the dif-
 100 ference between \bar{t} and the expected value of \hat{t} decays towards zero
 101 as the sequence length grows.

102 Notice also that when $f = g$ (i.e. the correction model matches
 103 the generating model) then $\bar{t} = t$. However, in general, \bar{t} need not
 104 be equal to t , except when $t = 0$. For example, if the generating
 105 model is the Jukes–Cantor model with some form of among-site
 106 rate heterogeneity and the correcting model that does not assume
 107 the same form of rate heterogeneity then \bar{t} can depend on t in a
 108 quite non-linear way (Soubrier et al., 2012).

109 In this paper we are interested in determining when the trans-
 110 formed evolutionary distances \bar{t} will favor a different tree to the
 111 tree generating the data. In particular, we explore an example of
 112 how ignoring the process of insertion and deletion (referred to
 113 jointly as indels hereafter) can lead to statistical inconsistency
 114 in an otherwise correctly modeled substitution process. Inconsistency
 115 occurs in this case even when the alignment of residues
 116 is correct.

117 Susko et al. (2004) studied general properties of \bar{t} as a function
 118 of t . If this function is linear (i.e. when the correction model
 119 matches the generating model up to a constant factor) then
 120 distance-based tree estimation will be statistically consistent. If
 121 the function is concave, inference can be inconsistent and posi-
 122 tively misleading due to long branch attraction. They also show
 123 that if the function is convex, two long branch repulsion trees
 124 are expected to be equally preferred over the correct tree. In
 125 Appendix A we establish a more general result: outside of the spe-
 126 cial case where the correcting generating model matches the gen-
 127 erating model up to a constant factor, there will always exist tree
 128 shapes for which neighbor-joining will estimate a single incorrect
 129 tree from \bar{t} . The tree shapes used to demonstrate this result are the
 130 familiar Felsenstein-zone tree (Fig. 1A; Felsenstein, 1978) and a
 131 tree that we refer to as the “twisted Farris-zone” tree (Fig. 1B).
 132 “Farris-zone” tree is used to refer to tree shapes that exhibit long
 133 branch repulsion under certain conditions of model violation, and
 134 this asymmetrical (“twisted”) variant has branch lengths which
 135 will result in a single incorrect tree topology being preferred if
 136 the corrected distance function is convex.

137 2.1. The gaps as missing data convention

138 It is common practice to treat a gap in a sequence as missing
 139 data in phylogenetic estimation based on distances, parsimony
 140 scores or likelihoods. In the context of a pairwise distance

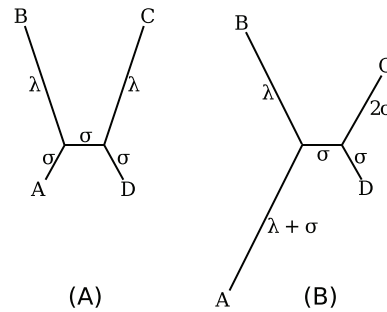


Fig. 1. (A) The Felsenstein-zone tree with branch lengths used in the proof of Lemma 3; (B) The “twisted Farris-zone” tree used in the proof of Lemma 4.

141 calculation, this treatment means that positions with a gap in
 142 either sequence are disregarded because they cannot be counted
 143 as either a similarity or a difference. Omitting indels from distance
 144 corrections obviously forfeits the opportunity for learning about
 145 the evolutionary distance from insertions and deletion events.
 146 However, one may hope that treating sites with gaps as missing
 147 data would not perturb a distance estimate that relies solely on
 148 substitution events. If the substitution and indel processes are
 149 completely independent, and have the same stationary nucleotide
 150 frequencies, this is the case.

151 Consider the case of sequences that are generated by: a
 152 time-reversible stochastic process of insertions and deletion, and
 153 a model of substitutions for which there is a statistically consistent
 154 distance correction. If the alignment is known without error, then
 155 the only effect of the indel process is to introduce a fraction of sites,
 156 z , for which one sequence lacks a residue and the other sequence
 157 has a residue. These are the gapped positions in a pairwise align-
 158 ment. Note that the presence of gap in a column in the alignment
 159 is not handled by deleting the column. The gap only affects pair-
 160 wise comparisons involving a sequence which contains a gap. A full
 161 description for z for a full alignment would require some additional
 162 notation to indicate which sequences are being compared. Our
 163 argument below applies to any pairwise distance, so we simply
 164 use $z(t)$ to describe the expected proportion of gapped position
 165 in any pairwise distance for sequences separated by path length, t .

166 The fraction of gapped positions will be a function of the evolu-
 167 tionary distance with: $z(0) = 0$ because at no distance there are no
 168 opportunities for indels, and $z(t) < 1$ for all t . The latter property
 169 can be seen by treating one of the two sequences as if it were
 170 the ancestral sequence. This is permissible because we have
 171 assumed that the indel process is time reversible. The probability
 172 of a residue surviving from the ancestral sequence to the descen-
 173 dant sequence is described by an exponential function with rate
 174 parameter controlled by the rate of deletions. This probability
 175 remains > 0 for all values of the evolutionary distance, hence there
 176 is a non-zero probability of an ungapped position, and $z(t)$ cannot
 177 equal 1.

178 In a typical consistency proof, we consider sequences of ever
 179 increasing length. We note that indel models (e.g. the TKF91
 180 model; Thorne et al. (1991)) imply an equilibrium sequence length.
 181 Here we discuss statistical consistency by considering what hap-
 182 pens as the number of loci increases without bound, but the equi-
 183 librium length of each locus is determined by the parameters of the
 184 indel model. Hence the total sequence length approaches infinity,
 185 while it is still appropriate to describe the sequence as being gen-
 186 erated by the indel process.

187 For the standard substitution models, we can consistently esti-
 188 mate the distance if the indel process has insertion and deletion
 189 rates of 0. In this case there are no gapped columns and $z(t) = 0$.
 190 In the more general case, if we only consider site patterns in which
 191 no gaps occur, the probability of a site pattern s for branch length t

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