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(1)  $C_n(\mathbf{x}) = 0$  if **x** is such that  $x_i = 0$  for some  $i \in \{1, 2, ..., n\}$ .

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J.J. Arias-García et al. / Fuzzy Sets and Systems ••• (••••) •••-•••

(2) 
$$C_n(\mathbf{x}) = x_i$$
 if  $\mathbf{x}$  is such that  $x_j = 1$  for all  $i \neq j$ .  
(3)  $C_n$  is *n*-increasing, i.e., for any *n*-box  $\mathbf{P} = \sum_{i=1}^n [a_i, b_i] \subseteq [0, 1]^n$  it holds that  
 $V_{C_n}(\mathbf{P}) = \sum_{\mathbf{x} \in \text{vertices}} (-1)^{S(\mathbf{x})} C_n(\mathbf{x}) \ge 0$ ,  
where  $S(\mathbf{x}) = \#\{i \in \{1, 2, ..., n\} \mid x_i = a_i\}$ .  
 $V_{C_n}(P)$  is called the  $C_n$ -volume of  $P$ .  
**Remark 1.** When  $n - k$  of the arguments of an *n*-copula are set equal to 1, we obtain a *k*-dimensional marginal of the  
*n*-copula, which itself is a *k*-copula.  
For a given *n*-copula  $C_n$ ,  $\hat{C}_n$  denotes the associated survival *n*-copula, and is given by  
 $\hat{C}_n(x_1, ..., x_n) = \sum_{i=1}^n x_i - (n-1) + \sum_{i < j}^n C_n(1, ..., 1 - x_i, ..., 1 - x_j, ..., 1)$   
 $-\sum_{i < j < k}^n C_n(1, ..., 1 - x_i, ..., 1 - x_j, ..., 1 - x_k, ..., 1) + ...$   
 $+ (-1)^n C_n(1 - x_1, 1 - x_2, ..., 1 - x_n)$ . (1)  
Survival copulas have a clear probabilistic interpretation: if the random vector  $(U_1, ..., U_n)$  has the copula  $C_n$  as its  
joint distribution function, then  $\hat{C}_n$  is the joint distribution function of the random vector  $(1 - U_1, ..., 1 - U_n)$ . This  
probabilistic interpretation has led to several studies of the transformations were studied in [7,9] for bivariate aggregation  
functions, while a specific transformation was studied in [11] for higher dimensions.  
Due to Sklar's theorem, which states that any continuous multivariate distribution function can be written in terms  
of its *n* univariate marginals by means of a unique *n*-copula have become one of the most important tools  
for the study of certain types of properties of random vectors, such as stochastic dependence (see [14,35] for more

details on *n*-copulas).

As mentioned before, an example of a property that can be directly studied using copulas is the property of radial symmetry. It can be easily shown that a continuous random vector  $(X_1, ..., X_n)$  is radially symmetric about  $(a_1, ..., a_n)$  if and only if, for any  $j \in \{1, ..., n\}$ ,  $X_j - a_j$  has the same distribution as  $a_j - X_j$ , and  $C_n = \hat{C}_n$ , where  $C_n$  is the copula associated to the random vector  $(X_1, ..., X_n)$ . Due to this characterization, we say that an *n*-copula  $C_n$  is radially symmetric if it satisfies the identity  $C_n = \hat{C}_n$ . Clearly, radial symmetry of an *n*-copula implies the radial symmetry of its lower dimensional marginals, however, the converse statement is not true.

It is important to remark that the concept of radial symmetry of an *n*-copula is different from the concept of symmetry. Symmetric copulas are the copulas associated to exchangeable random variables. Formally, we say that an *n*-copula  $C_n$  is symmetric if for any permutation  $\sigma$  of  $\{1, 2, ..., n\}$  and for any  $\mathbf{x} \in [0, 1]^n$  it holds that

$$C_n(x_1, x_2, \ldots, x_n) = C_n(x_{\sigma(1)}, x_{\sigma(2)}, \ldots, x_{\sigma(n)}).$$

Note that if an *n*-copula is symmetric, all of its *k*-dimensional marginals coincide for any  $k \in \{2, ..., n-1\}$ .

Radially symmetric copulas have a particular importance in stochastic simulation and statistics, as they can be
 used, in certain situations, in the multivariate version of the antithetic variates method, which is a variance reduction
 technique used in Monte Carlo methods [29]. Additionally, there has also been a growing interest in developing
 statistical tests for testing the presence of radial symmetry [1,3,10,18,22,38,41].

In the bivariate case, well-known examples of families of copulas that are radially symmetric are the Frank family and the Farlie–Gumbel–Morgenstern (FGM) family [35]. However, there are only a few families of *n*-copulas that are radially symmetric for  $n \ge 3$ , elliptical copulas being the best known [29]. Download English Version:

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