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COMPLITATIONAL

STATISTICS & DATA ANALYSIS

Computational Statistics and Data Analysis xx (xxxx) xxx-xxx

Contents lists available at ScienceDirect



Computational Statistics and Data Analysis

journal homepage: www.elsevier.com/locate/csda

A Gini-based unit root test

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HIGHLIGHTS

- A Gini-based unit root test is developed.
- The test relies on the semi-parametric Gini regression.
- Includes an in-depth numerical comparison of the Gini-based test and existing tests.
- Simulations indicate the superiority of the Gini-based test in some design settings.

ARTICLE INFO

Article history: Received 23 January 2014 Received in revised form 12 August 2014 Accepted 16 August 2014 Available online xxxx

Keywords: Time series analysis Unit root tests Gini regression Bootstrap

ABSTRACT

A Gini-based statistical test for a unit root is suggested. This test is based on the well-known Dickey–Fuller test, where the ordinary least squares (OLS) regression is replaced by the semi-parametric Gini regression in modeling the AR process. A residual-based bootstrap is used to find critical values. The Gini methodology is a rank-based methodology that takes into account both the variate values and the ranks. Therefore, it provides robust estimators that are rank-based, while avoiding loss of information. Furthermore, the Gini methodology relies on first-order moment assumptions, which validates its use for a wide range of distributions. Simulation results validate the Gini-based test and indicate its superiority in some design settings in comparison to other available procedures. The Gini-based test opens the door for further developments such as a Gini-based cointegration test.

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1. Introduction

In most of the literature dealing with time series analysis, underlying dependencies of the time series are modeled based on variance and covariance as measures of variability and association, respectively. This research develops a unit root test that is based on the Gini Mean Difference (hereafter GMD) as an alternative index of variability. The GMD index shares many properties of the variance, but the former can be more appropriate for distributions that depart from normality or symmetry. This measure is less sensitive to extreme observations than the variance because it takes into account both the values of the random variable and its ranks. In addition, the GMD is defined solely under first-order moment assumptions. To clarify the notations, we distinguish between population parameters and estimators by using upper-case letters in the population version and lower-case letters in the sample version.

1.1. Autoregressive unit root tests

We refer to a first-order univariate autoregression, denoted by AR(1), which satisfies

$$Y_t = \phi_0 + \phi_1 Y_{t-1} + \varepsilon_t,$$

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http://dx.doi.org/10.1016/j.csda.2014.08.012 0167-9473/© 2014 Published by Elsevier B.V.

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where ϕ_0 is the constant of the model, ϕ_1 is the parameter of the model and ε_t is an independent and identically distributed (i.i.d.) innovation process. If $\phi_1 = 1$, then Y_t is nonstationary. Testing for stationarity by detecting a unit root is an important task in the analysis and modeling of time series. Dickey and Fuller (1979) developed a procedure to test for the presence of a unit root (hereafter referred to as the DF test). The main objective of the DF test is to determine whether $H_0: \phi_1 = 1$ or, alternatively, H_1 : $\phi_1 < 1$.

The distribution of the appropriate *t*-statistic, based on applying the OLS estimator for the nonstationarity parameter, 6 is nonstandard and cannot be analytically evaluated. Dickey and Fuller (1979) used the Monte Carlo simulation method to 7 03 tabulate the percentiles of the DF t-statistic distribution based on $\varepsilon_t \sim i.i.d N(0, \sigma_s^2)$ innovations. The DF t-statistic is 8

$$DF_{t-stat} = \frac{\hat{\phi}_1^{OLS} - 1}{\hat{S}D(\hat{\phi}_1^{OLS})},$$

where $\hat{\phi}_1^{OLS} = \frac{\operatorname{cov}(Y_t, Y_{t-1})}{\operatorname{cov}(Y_{t-1}, Y_{t-1})}$ is the OLS estimator for ϕ_1 , $\hat{S}D(\hat{\phi}_1^{OLS}) = \left(\hat{\sigma}_{\varepsilon}^2 \left(\sum_{t=1}^n Y_{t-1}^2\right)^{-1}\right)^{0.5}$ and $\hat{\sigma}_{\varepsilon}^2$ is the least squares estimator 11

Extensive attempts to improve this test and to find alternative or superior tests followed. Leybourne (1995) suggested using the maximum of DF t-statistics based on applying the OLS regressions twice, looking both forward and backward at the series. The critical values for the test are obtained in a manner similar to that performed in the original DF test discussed above using Monte-Carlo simulation. Elliott et al. (1996) and Ng and Perron (2001) suggested a class of unit root tests that are based on generalized least squares detrending of the series and then applying the DF test on the detrended data.

An important development in unit root tests includes the use of resampling methods for calculating critical values; see, for 17 example, a survey of such methods in Palm et al. (2008). A first residual-based bootstrap version of the DF test was proposed 18 by Ferretti and Romo (1996). Later, Moreno and Romo (2000) used a bootstrap procedure based on the LAD estimator, and 19 more recently. Moreno and Romo (2012) suggested a family of unit root bootstrap tests for the infinite variance case. 20

Another important development appears in Muller and Elliott (2003), who revealed that the initial condition (Y_0) has a 21 non-negligible influence on the finite sample performances of unit root tests. In general, it is difficult to rule out, a priori, the 22 existence of small or large values of Y_0 . Elliott and Müller (2006) suggest statistics whose power is less sensitive to the size 23 of Y_0 , Harvey and Leybourne (2005) and Harvey et al. (2009) recommend a union of test rejection decision rules to improve 24 the performances of the statistical procedure. 25

Hallin et al. (2011) propose a class of distribution-free rank-based tests for the null hypothesis of a unit root that takes into 26 account several initial values for the series. They use three test statistics that are based on a choice of a reference density func-27 tion, which need not be the unknown actual density of the innovations. The first test statistic is based on the Gaussian refer-28 ence density and is defined as $T_{vdW}^{(n)} = \frac{1}{\sqrt{n}} \sum_{t=1}^{n} \left(\frac{t}{n+1} - \frac{1}{2}\right) \Phi^{-1}\left(\frac{R_t}{n+1}\right)$, where R_t are the ranks of the increments $\Delta Y_t = Y_t - Y_{t-1}$ and Φ denotes the standard normal distribution. This statistic is also known as the normal or van der Waerden score. The 29 30 second test statistic is based on the double-exponential distribution (Laplace or sign test scores), $T_L^{(n)} = \frac{1}{\sqrt{n}} \sum_{t=1}^n \left(\frac{t}{n+1} - \frac{1}{2}\right)$ 31 sign $\left(\frac{R_t}{n+1} - \frac{1}{2}\right)$. The third test statistic is based on the logistic distribution (Wilcoxon scores), $T_W^{(n)} = \frac{\pi}{\sqrt{3n}} \sum_{t=1}^n \left(\frac{t}{n+1} - \frac{1}{2}\right)$ $\left(\left(1 - \frac{n+1-R_t}{R_t}\right) / \left(1 + \frac{n+1-R_t}{R_t}\right)\right)$. These three test statistics are denoted here as HAW-vdW, HAW-Laplace and HAW-Wilcoxon, respectively. Asymptotic results and simulated quantiles for these statistics for several sample sizes are given 32 33 34

in Hallin et al. (2011). The results for finite samples indicate that for a broad range of non-zero initial values and for a variety 35 of heavy-tailed innovation densities, the suggested rank-based tests outperform a variety of unit root tests. 36

Rank-based tests, which intrinsically involve some loss of information of the real values, are expected to perform well 37 under heavy-tailed distributions. In this paper, we propose a Gini-based unit root test that relies on both the real values and 38 the ranks, while avoiding loss of information. 39

1.2. The Gini methodology 40

The GMD is an alternative index of variability that is used in this research instead of the variance. The most prevalent 41 presentation of the GMD index is the expected absolute difference between two independent and identically distributed 42 (i.i.d.) variables X_1 and X_2 (Gini, 1914). Formally, the GMD of X is defined as 43

$$G_X = E |X_1 - X_2|.$$
 (3)

Alternatively, the GMD can be expressed as a special case of a covariance, i.e., four times the covariance of X, a random 45 variable, and $F_X(X)$, its cumulative distribution function (Lerman and Yitzhaki, 1984). Formally, 46

$$G_{X} = 4COV(X, F_{X}(X)),$$

(4)

where $F_X(X)$ is the cumulative distribution function of X.

The GMD index shares many properties of the variance, but the former can be more informative for distributions that 49 depart from normality or symmetry. Both measures are based on weighted averages of the distances between each pair of 50 i.i.d. variables. The fundamental difference is the method used to measure the distance. The GMD distance function is referred 51 to as the "city block" distance, which allows one to move only in the vertical and horizontal directions. The variance distance 52 function is Euclidean, which allows one to move in any desired direction. 53

Please cite this article in press as: Shelef, A., A Gini-based unit root test. Computational Statistics and Data Analysis (2014), http://dx.doi.org/10.1016/j.csda.2014.08.012

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