



# Some integrals involving multivariate Hermite polynomials: Application to evaluating higher-order local powers

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## ABSTRACT

We present the formula for a certain integral with respect to multivariate Hermite polynomials. Such integrals are used for deriving higher-order local power functions of asymptotically chi-squared tests. As an example, we provide asymptotic expansion for the local power function of Rao's score test.

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## 1. Introduction

Let  $\phi_{\Sigma}(\cdot)$  be the density of the  $p$ -variate normal distribution  $N_p(\mathbf{0}_p, \Sigma)$ ;

$$\phi_{\Sigma}(\mathbf{v}) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \exp\left(-\frac{\mathbf{v}' \Sigma^{-1} \mathbf{v}}{2}\right), \quad \mathbf{v} = (v_1, \dots, v_p)',$$

where  $\Sigma = [\sigma_{j,j'}]_{j,j' \in \{1, \dots, p\}}$  is a  $p \times p$  positive definite matrix. For any  $n \in \mathbf{N}$ , we set

$$\phi_{\Sigma}^{j_1, \dots, j_n}(\mathbf{v}) = (-1)^n \left( \prod_{i=1}^n \frac{\partial}{\partial v_{j_i}} \right) \phi_{\Sigma}(\mathbf{v}), \quad j_1, \dots, j_n \in \{1, \dots, p\};$$

$\phi_{\Sigma}^{j_1, \dots, j_n}(\mathbf{v}) / \phi_{\Sigma}(\mathbf{v})$  is referred to as the multivariate Hermite polynomial. For example,

$$\begin{aligned} \phi_{\Sigma}^{j_1}(\mathbf{v}) &= \tilde{v}_{j_1} \phi_{\Sigma}(\mathbf{v}), \\ \phi_{\Sigma}^{j_1, j_2}(\mathbf{v}) &= (\tilde{v}_{j_1} \tilde{v}_{j_2} - \sigma^{j_1, j_2}) \phi_{\Sigma}(\mathbf{v}), \\ \phi_{\Sigma}^{j_1, j_2, j_3}(\mathbf{v}) &= (\tilde{v}_{j_1} \tilde{v}_{j_2} \tilde{v}_{j_3} - \langle 3 \rangle \sigma^{j_1, j_2} \tilde{v}_{j_3}) \phi_{\Sigma}(\mathbf{v}), \end{aligned}$$

where  $\tilde{v}_j = \sigma^{j, j'} v_{j'}$  (we used the standard summation convention instead of  $\sum_{j'=1}^p \sigma^{j, j'} v_{j'}$ ), with  $\sigma^{j, j'}$  being the  $(j, j')$ th element of  $\Sigma^{-1}$ . Here, the notation  $\langle n \rangle$  before a term with indices is a sum of  $n$  similar terms obtained by index permutation, e.g.,  $\langle 3 \rangle \sigma^{j_1, j_2} \tilde{v}_{j_3} = \sigma^{j_1, j_2} \tilde{v}_{j_3} + \sigma^{j_1, j_3} \tilde{v}_{j_2} + \sigma^{j_2, j_3} \tilde{v}_{j_1}$ . Note that

$$\phi_{\Sigma}^{j_1, \dots, j_n}(\mathbf{v}) = \sum_{i=0}^{\lfloor n/2 \rfloor} \left\langle \binom{n}{2i} \right\rangle_{j_1 \dots j_{2i} j_{2i+1} \dots j_n} \left[ (-1)^i \left\langle \frac{(2i)!}{2^i i!} \right\rangle_{\text{all } i \text{ pairs}} \sigma^{j_1, j_2} \dots \sigma^{j_{2i-1}, j_{2i}} \right] \tilde{v}_{j_{2i+1}} \dots \tilde{v}_{j_n} \phi_{\Sigma}(\mathbf{v})$$

(e.g., Barndorff-Nielsen and Cox, 1989, page 151, with somewhat different notation).

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Given  $\boldsymbol{\tau} = (\tau_1, \dots, \tau_p)'$ , it is well known (e.g., [Anderson, 2003](#), page 82) that if  $\mathbf{Z} \sim N_p(\boldsymbol{\tau}, \boldsymbol{\Sigma})$ , then,  $\mathbf{Z}'\boldsymbol{\Sigma}^{-1}\mathbf{Z}$  has a noncentral  $\chi^2$ -distribution with  $p$  degrees of freedom and noncentral parameter  $\boldsymbol{\tau}'\boldsymbol{\Sigma}^{-1}\boldsymbol{\tau}$ ;

$$\Pr[\mathbf{Z}'\boldsymbol{\Sigma}^{-1}\mathbf{Z} \leq x] = \int_0^x g_p(t; \boldsymbol{\tau}'\boldsymbol{\Sigma}^{-1}\boldsymbol{\tau}) dt \quad \text{for } x > 0, \quad (1)$$

where

$$g_p(t; \omega^2) = \sum_{\ell=0}^{\infty} \frac{(\omega^2/2)^\ell \exp(-\omega^2/2)}{\ell!} \frac{t^{(v+2\ell)/2-1} \exp(-t/2)}{2^{(v+2\ell)/2} \Gamma((v+2\ell)/2)}$$

is the density of a noncentral  $\chi^2$ -distribution with  $v$  degrees of freedom and noncentral parameter  $\omega^2$ . As usual, we write  $G_v(x; \omega^2) = \int_0^x g_v(t; \omega^2) dt$  and  $G_v^-(x; \omega^2) = 1 - G_v(x; \omega^2)$ .

We are concerned with the integral of  $\phi_{\boldsymbol{\Sigma}}^{j_1, \dots, j_q}(\mathbf{v})$  over the convex set  $\{\mathbf{v} \in \mathbf{R}^p : (\mathbf{v} + \boldsymbol{\tau})'\boldsymbol{\Sigma}^{-1}(\mathbf{v} + \boldsymbol{\tau}) \leq x\}$ ;

$$\int_{(\mathbf{v}+\boldsymbol{\tau})'\boldsymbol{\Sigma}^{-1}(\mathbf{v}+\boldsymbol{\tau}) \leq x} \phi_{\boldsymbol{\Sigma}}^{j_1, \dots, j_q}(\mathbf{v}) d\mathbf{v} = - \int_{\mathcal{D}_{\boldsymbol{\Sigma}}(x)} \phi_{\boldsymbol{\Sigma}}^{j_1, \dots, j_q}(\mathbf{v}) d\mathbf{v} \quad \text{for } q \in \mathbf{N}, \quad (2)$$

where  $\mathcal{D}_{\boldsymbol{\Sigma}}(x) = \{\mathbf{v} \in \mathbf{R}^p : (\mathbf{v} + \boldsymbol{\tau})'\boldsymbol{\Sigma}^{-1}(\mathbf{v} + \boldsymbol{\tau}) > x\}$  (we used  $\int_{\mathbf{R}^p} \phi_{\boldsymbol{\Sigma}}^{j_1, \dots, j_q}(\mathbf{v}) d\mathbf{v} = 0$ ; the orthogonality of the multivariate Hermite polynomial). Such integrals (2) for  $q = 1, 2, 3, 4, 5, 6$  are implicitly used in the derivation of asymptotic expansions for the distributions of asymptotically chi-squared test statistics (e.g., [Mukerjee, 1990a,b](#), [Fujikoshi, 1997](#), [Bravo, 2003, 2004](#), and [Kakizawa, 2010a,b, 2012a,b, 2013](#)). More precisely, we suppose that a random vector  $\mathbf{U}^{(N)} = (U_1^{(N)}, \dots, U_p^{(N)})' \xrightarrow{d} N_p(\boldsymbol{\tau}, \boldsymbol{\Sigma})$  possesses cumulants

$$\begin{aligned} \text{cum}(U_j^{(N)}) &= \tau_j + \sum_{\ell \geq 1} N^{-\ell/2} c_j^{(\ell)}, \\ \text{cum}(U_{j_1}^{(N)}, U_{j_2}^{(N)}) &= \sigma_{j_1 j_2} + \sum_{\ell \geq 1} N^{-\ell/2} c_{j_1 j_2}^{(\ell)}, \\ \text{cum}(U_{j_1}^{(N)}, \dots, U_{j_s}^{(N)}) &= \sum_{\ell \geq s-2} N^{-\ell/2} c_{j_1 \dots j_s}^{(\ell)}, \quad s \in \{3, 4, \dots\}, \end{aligned}$$

where  $c_{\#}^{(\ell)}$ 's are the coefficients of  $N^{-\ell/2}$  in the cumulants. We write  $e_{\ell}(\mathbf{z}) = \sum_{r=1}^{\ell+2} \frac{1}{r!} c_{j_1 \dots j_r}^{(\ell)} z_{j_1} \dots z_{j_r}$  for  $\mathbf{z} = (z_1, \dots, z_p)' \in \mathbf{C}^p$ , and prepare the ordinary partial Bell polynomial  $B_{\ell, k}^{\circ} = B_{\ell, k}^{\circ}(x_1, x_2, \dots)$  for positive integers  $\ell$  and  $k$  ( $\ell \geq k$ ), defined as the formal series expansion  $(\sum_{\ell \geq 1} x_{\ell} y^{\ell})^k = \sum_{\ell \geq k} B_{\ell, k}^{\circ} y^{\ell}$ , i.e.,  $B_{\ell, k}^{\circ} = \sum \frac{\ell!}{i_1! i_2! \dots} x_1^{i_1} x_2^{i_2} \dots$  (see [Comtet, 1974](#), page 136), where the summation takes over all nonnegative integers  $i_1, i_2, \dots$ , such that  $i_1 + i_2 + i_3 + \dots = \ell$  and  $i_1 + 2i_2 + 3i_3 + \dots = k$  (e.g.,  $B_{1,1}^{\circ} = x_1$ ,  $B_{2,1}^{\circ} = x_2$ ,  $B_{2,2}^{\circ} = x_1^2$ ). Then, as in [Taniguchi \(1991, page 14\)](#) (see also [Withers and Nadarajah, 2010](#)), the characteristic function of  $\mathbf{U}^{(N)}$  is formally expanded as

$$\begin{aligned} & \exp\left\{\tau_j(it_j) + \frac{1}{2!} \sigma_{j_1 j_2}(it_{j_1})(it_{j_2}) + \sum_{s \geq 1} \frac{1}{s!} \sum_{\ell \geq \max(1, s-2)} N^{-\ell/2} c_{j_1 \dots j_s}^{(\ell)}(it_{j_1}) \dots (it_{j_s})\right\} \\ &= \exp\left\{it_j \tau_j - \frac{1}{2} \sigma_{j_1 j_2} t_{j_1} t_{j_2} + \sum_{\ell \geq 1} N^{-\ell/2} e_{\ell}(\mathbf{it})\right\} \\ &= \exp\left(it_j \tau_j - \frac{1}{2} \sigma_{j_1 j_2} t_{j_1} t_{j_2}\right) \left[1 + \sum_{k \geq 1} \frac{1}{k!} \sum_{\ell \geq k} N^{-\ell/2} B_{\ell, k}^{\circ}(e_1(\mathbf{it}), e_2(\mathbf{it}), \dots)\right] \\ &= \exp\left(it_j \tau_j - \frac{1}{2} \sigma_{j_1 j_2} t_{j_1} t_{j_2}\right) \left[1 + \sum_{\ell \geq 1} N^{-\ell/2} \sum_{k=1}^{\ell} \frac{1}{k!} B_{\ell, k}^{\circ}(e_1(\mathbf{it}), e_2(\mathbf{it}), \dots)\right] \\ &= \exp\left(it_j \tau_j - \frac{1}{2} \sigma_{j_1 j_2} t_{j_1} t_{j_2}\right) \left[1 + \sum_{\ell \geq 1} N^{-\ell/2} C_{\ell}^{\circ}(\mathbf{it})\right] \quad (\mathbf{i} = \sqrt{-1} \text{ and } \mathbf{t} = (t_1, \dots, t_p)' \in \mathbf{R}^p), \end{aligned}$$

where  $C_{\ell}^{\circ}(\mathbf{z}) = \sum_{k=1}^{\ell} \frac{1}{k!} B_{\ell, k}^{\circ}(e_1(\mathbf{z}), e_2(\mathbf{z}), \dots)$  is a polynomial of degree  $3\ell$  (with  $C_{\ell}^{\circ}(\mathbf{0}_p) = 0$ ); note that  $C_1^{\circ}(\mathbf{z}) = e_1(\mathbf{z})$  and  $C_2^{\circ}(\mathbf{z}) = e_2(\mathbf{z}) + e_1^2(\mathbf{z})/2$ . It follows that the distribution of  $\mathbf{U}^{(N)}$  over a certain Borel set  $A \subset \mathbf{R}^p$  may be approximated as

$$\Pr[\mathbf{U}^{(N)} \in A] = \int_A \left[1 + \sum_{\ell=1}^{r-2} N^{-\ell/2} C_{\ell}^{\circ}\left(-\frac{\partial}{\partial v_{j_1}}, \dots, -\frac{\partial}{\partial v_{j_p}}\right)\right] \phi_{\boldsymbol{\Sigma}}(\mathbf{v} - \boldsymbol{\tau}) d\mathbf{v} + o(N^{-(r-2)/2}).$$

Especially, the choice  $A = \{\mathbf{v} \in \mathbf{R}^p : \mathbf{v}'\boldsymbol{\Sigma}^{-1}\mathbf{v} \leq x\}$  yields

$$\Pr[(\mathbf{U}^{(N)})'\boldsymbol{\Sigma}^{-1}\mathbf{U}^{(N)} \leq x] = G_p(x; \boldsymbol{\tau}'\boldsymbol{\Sigma}^{-1}\boldsymbol{\tau}) - \sum_{\ell=1}^{r-2} N^{-\ell/2} \int_{\mathcal{D}_{\boldsymbol{\Sigma}}(x)} C_{\ell}^{\circ}\left(-\frac{\partial}{\partial v_{j_1}}, \dots, -\frac{\partial}{\partial v_{j_p}}\right) \phi_{\boldsymbol{\Sigma}}(\mathbf{v}) d\mathbf{v} + o(N^{-(r-2)/2}). \quad (3)$$

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