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## Rate of Poisson approximation for nearest neighbor counts in large-dimensional Poisson point processes

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### ABSTRACT

Consider two independent homogeneous Poisson point processes  $\Pi$  of intensity  $\lambda$  and  $\Pi'$  of intensity  $\lambda'$  in  $d$ -dimensional Euclidean space. Let  $q_{k,d}$ ,  $k = 0, 1, \dots$ , be the fraction of  $\Pi$ -points which are the nearest  $\Pi$ -neighbor of precisely  $k$   $\Pi'$ -points. It is known that as  $d \rightarrow \infty$ , the  $q_{k,d}$  converge to the Poisson probabilities  $e^{-\lambda'/\lambda} (\lambda'/\lambda)^k / k!$ ,  $k = 0, 1, \dots$ . We derive the (sharp) rate of convergence  $d^{-1/2} (4/3\sqrt{3})^d$ , which is related to the asymptotic behavior of the variance of the volume of the typical cell of the Poisson–Voronoi tessellation generated by  $\Pi$ . An extension to the case involving more than two independent Poisson point processes is also considered.

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In this note, we consider two independent homogeneous Poisson point processes  $\Pi$  of intensity  $\lambda$  and  $\Pi'$  of intensity  $\lambda'$  in  $d$ -dimensional Euclidean space  $\mathbb{R}^d$ . Let  $p_{k,d}$ ,  $k = 0, 1, \dots$ , be the fraction of  $\Pi$ -points which are the nearest  $\Pi$ -neighbor of precisely  $k$  other  $\Pi$ -points, and  $q_{k,d}$ ,  $k = 0, 1, \dots$ , the fraction of  $\Pi$ -points which are the nearest  $\Pi$ -neighbor of precisely  $k$   $\Pi'$ -points. Here for given  $\mathbf{v} \in \mathbb{R}^d$ , a point  $Q \in \Pi$  is called the nearest  $\Pi$ -neighbor of  $\mathbf{v}$  if  $\|Q - \mathbf{v}\|_d < \|\mathbf{u} - \mathbf{v}\|_d$  for all  $\mathbf{u} \in \Pi \setminus \{Q\}$  where  $\|\cdot\|_d$  denotes the Euclidean norm in  $\mathbb{R}^d$ . By ergodic-type arguments,  $p_{k,d}$  and  $q_{k,d}$  are well defined. In their Theorems 5 and 10, Newman et al. (1983) proved that

$$\lim_{d \rightarrow \infty} p_{k,d} = e^{-1}/k! \quad \text{and} \quad \lim_{d \rightarrow \infty} q_{k,d} = e^{-\rho} \rho^k / k!, \quad (1)$$

where  $\rho = \lambda'/\lambda$ . (See also Newman and Rinott, 1985.)

The limit results in (1) can also be formulated in terms of a “typical point”  $Q$  of  $\Pi$ , to be translated so that  $Q = \mathbf{0} = (0, \dots, 0)$ , the origin in  $\mathbb{R}^d$ . We will refer to  $\mathbf{0}$  as the typical point of  $\Pi$ . (Note that since  $\Pi$  is a Poisson point process, its Palm distribution at  $\mathbf{0}$  is equivalent to the distribution of  $\Pi$  with an independently added point at  $\mathbf{0}$ ; see e.g. Daley and Vere-Jones (2007, Proposition 13.1.VII).) Let  $M_d$  be the number of  $\Pi$ -points which have  $\mathbf{0}$  as their nearest  $\Pi$ -neighbor, and  $N_{\rho,d}$  the number of  $\Pi'$ -points which have  $\mathbf{0}$  as their nearest  $\Pi$ -neighbor, i.e.

$$\begin{aligned} M_d &= \#\{\mathbf{u} \in \Pi : \|\mathbf{u} - \mathbf{0}\|_d < \|\mathbf{u} - \mathbf{v}\|_d \text{ for all } \mathbf{v} \in \Pi \setminus \{\mathbf{u}\}\}, \\ N_{\rho,d} &= \#\{\mathbf{u}' \in \Pi' : \|\mathbf{u}' - \mathbf{0}\|_d < \|\mathbf{u}' - \mathbf{v}\|_d \text{ for all } \mathbf{v} \in \Pi\} \\ &= \#(\Pi' \cap C_d), \end{aligned} \quad (2)$$

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where

$$C_d = \{\mathbf{x} \in \mathbb{R}^d : \|\mathbf{x} - \mathbf{0}\|_d < \|\mathbf{x} - \mathbf{u}\|_d \text{ for all } \mathbf{u} \in \Pi\}, \quad (3)$$

the (typical) Voronoi cell centered at  $\mathbf{0}$  generated by  $\Pi \cup \{\mathbf{0}\}$  (cf. Okabe et al., 2000). Note that  $\mathcal{L}(M_d)$ , the distribution of  $M_d$ , is independent of  $\lambda$  and  $\lambda'$ , and  $\mathcal{L}(N_{\rho,d})$  depends on  $\lambda$  and  $\lambda'$  only through  $\rho = \lambda'/\lambda$ . Then (1) is equivalent to the limit results that as  $d \rightarrow \infty$ ,  $M_d$  and  $N_{\rho,d}$  converge in distribution to  $Po(1)$  and  $Po(\rho)$ , respectively, where  $Po(\rho)$  denotes the Poisson distribution with mean  $\rho$ .

In the present note, we derive the rate of convergence for  $N_{\rho,d}$  as stated below.

**Theorem 1.** For any given  $\rho_0 > 0$ , there exists a constant  $c_1(\rho_0) > 0$  such that for all  $1 \leq d < \infty$  and  $0 < \rho \leq \rho_0$ ,

$$c_1(\rho_0)\rho^2 \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d \leq d_{TV}(\mathcal{L}(N_{\rho,d}), Po(\rho)) \leq c_2\rho(1 - e^{-\rho}) \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d, \quad (4)$$

where  $d_{TV}$  denotes the total variation distance and  $c_2 > 0$  is a constant (independent of  $\rho_0$ ).

**Proof.** Without loss of generality, assume  $\lambda = 1$ , so that  $\rho = \lambda'/\lambda = \lambda'$ . Note that  $N_{\rho,d}$  has a mixed Poisson distribution. By (2) and (3), the conditional distribution of  $N_{\rho,d}$  given  $\mu_d(C_d) = v$  is  $Po(\lambda'v) = Po(\rho v)$  where  $\mu_d(S)$  denotes the  $d$ -dimensional Lebesgue measure (volume) of measurable  $S \subset \mathbb{R}^d$ . Also,  $E[\mu_d(C_d)] = 1/\lambda = 1$ , and by Alishahi and Sharifitabar (2008, Theorem 3.1),

$$\text{var}(\mu_d(C_d)) \leq \frac{c_2}{\lambda^2} \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d = c_2 \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d \quad \text{for all } 1 \leq d < \infty,$$

for some constant  $c_2 > 0$ . By Barbour et al. (1992, Theorem 1.C(ii)), we have

$$\begin{aligned} d_{TV}(\mathcal{L}(N_{\rho,d}), Po(\rho)) &\leq \rho^{-1}(1 - e^{-\rho})\text{var}(\rho\mu_d(C_d)) \\ &\leq c_2\rho(1 - e^{-\rho}) \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d, \end{aligned}$$

establishing the upper bound.

To derive the lower bound, fix a (large)  $u > 0$  and consider the ball of volume  $u$  centered at  $\mathbf{0}$ , denoted by  $B = B_{u,d}$ . Applying Lemma 1 below with  $\alpha_1 = \rho\mu_d(C_d \cap B) \leq \rho u$ ,  $\alpha_2 = \rho\mu_d(C_d \setminus B)$ ,  $\beta_1 = E[\alpha_1] = \rho E[\mu_d(C_d \cap B)] \leq \rho u$ ,  $\beta_2 = E[\alpha_2] = \rho E[\mu_d(C_d \setminus B)] = \rho - \beta_1$ , we have

$$\begin{aligned} d_{TV}(\mathcal{L}(N_{\rho,d}), Po(\rho)) &\geq P(N_{\rho,d} = 0) - e^{-\rho} \\ &= E[e^{-\rho\mu_d(C_d)}] - e^{-\rho} \\ &= E\{e^{-\alpha_1 - \alpha_2} - e^{-\beta_1 - \beta_2}\} \\ &\geq E\{e^{-\beta_1 - \beta_2}[h_1(\rho u)(\alpha_1 - \beta_1)^2 - (\alpha_1 - \beta_1)] - e^{-\alpha_1 - \beta_2}(\alpha_2 - \beta_2)\} \\ &= e^{-\rho} h_1(\rho u) \text{var}(\alpha_1) - e^{-\beta_2} E[e^{-\alpha_1}(\alpha_2 - \beta_2)] \\ &\geq e^{-\rho} h_1(\rho u) \text{var}(\rho\mu_d(C_d \cap B)) \\ &\geq e^{-\rho} h_1(\rho u) \rho^2 c \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d, \end{aligned}$$

where the last two inequalities follow from Lemma 2 (with  $S_1 = B$  and  $S_2 = \mathbb{R}^d \setminus B$ ) and Lemma 3 (with  $u \geq u_0$  and  $c > 0$  and  $u_0 < \infty$  appearing in the statement of Lemma 3). Noting that  $h_1$  is nonincreasing, we have for all  $1 \leq d < \infty$  and  $0 < \rho \leq \rho_0$ ,

$$\begin{aligned} d_{TV}(\mathcal{L}(N_{\rho,d}), Po(\rho)) &\geq ce^{-\rho_0} h_1(\rho_0 u) \rho^2 \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d \\ &= c_1(\rho_0) \rho^2 \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d, \end{aligned}$$

where  $c_1(\rho_0) = ce^{-\rho_0} h_1(\rho_0 u)$ . The proof is complete.  $\square$

**Remark 1.** The lower and upper bounds in (4) may be expressed as

$$\begin{aligned} \inf_{1 \leq d < \infty, 0 < \rho \leq \rho_0} d_{TV}(\mathcal{L}(N_{\rho,d}), Po(\rho)) / \left[ \rho^2 \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d \right] &> 0, \\ \sup_{1 \leq d < \infty, \rho > 0} d_{TV}(\mathcal{L}(N_{\rho,d}), Po(\rho)) / \left[ \rho(1 - e^{-\rho}) \frac{1}{\sqrt{d}} \left( \frac{4}{3\sqrt{3}} \right)^d \right] &< \infty. \end{aligned}$$

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