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## Analysis of size and temperature effects in the ductile to brittle transition region of ferritic steels

### C. Berejnoi<sup>a,\*</sup>, J.E. Perez Ipiña<sup>b</sup>

<sup>a</sup> Facultad de Ingeniería, Universidad Nacional de Salta, Avda. Bolivia 5150, 4400 Salta, Argentina
<sup>b</sup> CONICET/Facultad de Ingeniería, Universidad Nacional del Comahue, Avda. Buenos Aires 1400, 8300 Neuquén, Argentina

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#### ABSTRACT

This paper presents an analysis of ferritic steels in the ductile-to-brittle transition region that includes the determination of the temperature reference of the Master Curve, which assumes a Weibull distribution with fixed threshold and shape parameters for compact specimens of one inch thickness. Some differences arise between the scale parameter and the median of the distribution calculated from these specimens and those converted from other sizes. The dependence with size and temperature of the parameters of a non-fixed three parameter Weibull distribution were also analyzed. The estimated threshold and shape parameters resulted clearly temperature dependent, and different from those stated in the Master Curve.

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#### 1. Introduction

The characterization of fracture resistance of ferritic steels in the ductile-to-brittle transition region is problematic due to scatter in results, as well as size and temperature dependences.

Ferritic steels are, as defined in ASTM E1921-13 [1], "typically carbon, low-alloy, and higher alloy grades. Typical microstructures are bainite, tempered bainite, tempered martensite, and ferrite and pearlite. All ferritic steels have body centered cubic crystal structures that display ductile-to-cleavage transition temperature fracture toughness characteristics."

The statistical treatment is mainly based on Weibull statistics, which has been used with two (2P-W) (Eq. (1)) or three parameter (3P-W) (Eq. (2)). The parameters to be determined in the 2P-W distribution are the shape parameter, also known as Weibull slope ( $b_K$  or  $b_J$ ), and the scale parameter ( $K_0$  or  $J_0$ ). For a 3P-W distribution, the threshold parameter ( $K_{min}$  or  $J_{min}$ ) is added.

$$P = 1 - \exp\left[-\left(\frac{Jc}{J_0}\right)^{b_j}\right] \quad P = 1 - \exp\left[-\left(\frac{K_{Jc}}{K_0}\right)^{b_k}\right] \tag{1}$$

$$P = 1 - \exp\left[-\left(\frac{Jc - J_{min}}{J_0 - J_{min}}\right)^{b_J}\right] \quad P = 1 - \exp\left[-\left(\frac{K_{Jc} - K_{min}}{K_0 - K_{min}}\right)^{b_K}\right]$$
(2)

Data expressed in terms of *K* are derived from  $J_C(K_{IC})$  using Eq. (3).

\* Corresponding author. Tel.: +54 387 4255420.

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E-mail addresses: berejnoi@unsa.edu.ar (C. Berejnoi), juan.perezipina@fain.uncoma.edu.ar (J.E. Perez Ipiña).

Nomenciature	
a	initial crack length
$b_0$	initial specimen remaining ligament
b	Weibull shape parameter estimated from <i>Ic</i> data
b <sub>κ</sub>	Weibull shape parameter estimated from $K_{IC}$ data
$b_{k-1T}$	$b_k$ for 1T-C(T) size (original and converted)
$b_{Ki}$	Weibull shape parameter derived from estimated $b_1$
r	number of non-censored tests
В	specimen thickness
$B_{1T}$	equivalent 1T-C(T) thickness
C(T)	compact tension specimen
Ε	elastic modulus
Jc	experimental J-integral at the onset of cleavage fracture
J <sub>max</sub>	maximum allowed $J_c$ value
Jmin	Weibull threshold parameter estimated from $J_C$ data
Jo	Weibull scale parameter estimated from $J_C$ data
K <sub>Jc</sub>	elastic-plastic equivalent stress intensity factor derived from the $J_C$ value
K <sub>Jc-1T</sub> V	equivalent $\prod_{i=1}^{N} C(i)$ value of $K_{J_c}$
$K_{Jc(i)}$	$I-UI$ value of $K_{J_c}$
K <sub>Jmax</sub> K	median of $K_{i}$ distribution
K <sub>med</sub>	Weibull threshold parameter
K <sub>min</sub> K	minimum experimental K.
$K_{min-exp}$ $K_{min}(I)$	K <sub>min</sub> derived from estimated I <sub>min</sub>
$K_{min}(K)$	$K_{min}$ derived from $K_{i}$ data
Kmin_1T	$K_{min}$ for 1T-C(T) size (original and converted)
$K_0$	Weibull scale parameter
$K_{0-1T}$	$K_0$ for 1T-C(T) size (original and converted)
$K_0(J)$	$K_0$ derived from estimated $J_0$
$K_0(K)$	$K_0$ estimated from $K_{lc}$ data
MC	Master Curve
Ν	total number of tests
Т	test temperature
$T_0$	reference temperature
W	specimen width
ξ	factor relating $b_J$ and $b_{Kj}$
v	Poisson's ratio

$$K_{Jc} = \sqrt{\frac{EJ_c}{(1-v^2)}}$$

Nomonelature

(3)

where *E* is the elastic modulus and v is the Poisson's ratio.

For instance, Landes and Shaffer [2], Iwadate et al. [3], Anderson et al. [4], Landes et al. [5], and Heerens et al. [6,7] made use of a 2P-W distribution based on  $J_C$  values, while Landes and McCabe [8], Neville and Knott [9], and Perez Ipiña et al. [10] based their analysis on the 3P-W distribution using  $J_C$  data. The use of such distributions based on K values was promoted by Wallin, with a 2P-W distribution [11], and later with a 3P-W distribution [12].

Besides the possibility of working with two or three parameters, and also with J or K data, some authors have proposed a fixed shape parameter with a given value: 2 when working with  $J_C$  [4–6,13] and 4 when working with  $K_{JC}$  [1,12,14].

Although the 2P-Weibull slopes in terms of *J* and *K* are related by  $b_K/b_J = 2$ , this relationship does not apply when the third parameter (threshold parameter) is introduced. It was shown in previous papers [15,16] that the slopes ratio is not 2, and it is given by the factor  $\xi = \frac{2K_0}{K_0 + K_{min}}$ . The slope  $b_{Kj}$ , converted from  $b_J$  using the factor  $\triangleright$ , and the  $b_K$  estimated from  $K_{Jc}$  values are not equals, but they are similar.

ASTM [1] has adopted the Master Curve method for the analysis of fracture toughness in this region. The temperature dependence of the 1T-C(T) median fracture toughness is based on an empirical equation calibrated at the  $T_0$  temperature that corresponds to a  $K_{med}$  = 100 MPa m<sup>0.5</sup> for this size, and it is determined assuming a Weibull distribution with  $K_{min}$  = 20 MPa m<sup>0.5</sup> and slope  $b_K$  = 4 for the scatter treatment.

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